Seroprevalence of COVID-19 among health workers in the Kathmandu valley, Nepal (SEVID-KaV): a longitudinal cohort study

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Seroprevalence of COVID-19 among health workers in the Kathmandu valley, Nepal (SEVID-KaV): a longitudinal cohort study

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Abstract

Coronavirus Disease 2019 (COVID-19) burden is often underestimated when relying on case-based incidence reports. Seroprevalence studies accurately estimate infectious disease burden by estimating the population that has developed antibodies following an infection. Sero-Epidemiology of COVID-19 in the Kathmandu valley (SEVID-KaV) is a longitudinal survey of hospital-based health workers in the Kathmandu valley. Between December 3-25, we sampled 800 health workers from 20 hospitals and administered a questionnaire eliciting COVID-19 related history and tested for COVID-19 IgG antibodies. We then used a probabilistic multilevel regression model with post-stratification to correct for test accuracy, the effect of hospital-based clustering, and to establish representativeness. 522 (65.2%) of the participants were female, 372 (46%) were between ages 18-29, and 7 (0.9%) were 60 or above. 287 (36%) of the participants were nurses. About 23% of the participants previously had a PCR positive infection. 321 (40.13%) individuals tested positive for COVID-19 antibodies. Adjusted for test accuracy and weighted by age, gender and occupation category, the seroprevalence was 38.17% (95% Credible Interval (CrI) 29.26%–47.82%). Posterior predictive hospital-wise seroprevalence ranged between 38.1% (95% CrI 30.7%–44.1%) and 40.5% (95% CrI 34.7%–47.0%).
Introduction

Population-based seroprevalence studies have been useful in quantifying the cumulative incidence of the coronavirus disease 2019 (COVID-19) epidemic. Nepal (population 30 million) reported its first infection of SARS-CoV-2 on January 24, 2020—the same week several countries including France, Vietnam, United States, and India reported their first infection. Since then, more than half a million people have been diagnosed with COVID-19 in Nepal, among which more than 7000 have died as of June 2021.\(^1\)

Nepal’s first wave of COVID-19 infections peaked in November 2020. The number of cases fell over the winter, with test positivity rates close to 1% from a high of about 25% around October 2020. Although seroprevalence studies suggest that the cumulative burden of SARS-CoV-2 infection is often several-folds greater than the reported case burden,\(^2\) an ongoing second wave that began in March 2021 suggests that a significant percentage of the population was still susceptible to infection after the first wave.

Seroprevalence studies are helpful in estimating the true extent of viral spread because they detect seroconversion (i.e. presence of antibodies) after an infection, even among those without clinical or laboratory evidence of active disease. Based on our current understanding, in almost all infected individuals, antibodies against SARS-CoV2 appear within 3 weeks of symptom onset.\(^6,7\) When administered to a representative sample of the population, seroprevalence studies can help assess the prevalence of SARS-CoV-2 antibodies as well as cumulative incidence of COVID-19. Seroprevalence studies among at-risk populations like health workers may be a leading indicator of infection burden in the community. In this report, we present results from wave 1 of the Sero-Epidemiology of COVID-19 in the Kathmandu Valley (SEVID-KaV) study, a longitudinal-cohort seroprevalence study among hospital based health workers in Kathmandu, Nepal.

Methods
Study design and population

SEVID-KaV study participants were chosen by means of a two-stage cluster-stratified random sampling method. In the first stage, we formed a sampling frame of all hospitals in the Kathmandu valley (population of about 3,000,000) with more than 100 staff-members (63 out of 74 hospitals, with about 25700 staff members), following which we selected 20 hospitals based on the Probability Proportionate to Size (PPS) method. In the second stage, we selected 40 staff members from each of the 20 hospitals based on simple random sampling (SRS) when possible, resulting in a sample size of 800 health workers. This sample size was expected to estimate the seroprevalence with a margin of error of less than 5%. PPS with SRS ensures that each health worker in the selected hospitals has the same probability of being sampled in the study.

Sampling was conducted between December 3 to 25, 2020. Hospital staff (clinical as well as administrative) above 18 years of age were eligible to participate. Staff names were obtained from hospital human resources departments and randomly ordered using a spread-sheet. Hospital staff were then telephoned in that random order until 40 participants could be recruited for the study. Six hospitals were unable to provide phone numbers for staff members. There, a convenience sample of 40 staff members was randomly sampled from among those present at the hospital on the day of sampling.

Ethical approval for all study sites was granted by the Nepal Health Research Council’s Ethical Review Board (ERB) (Approval reference number: 609). All except two hospitals that were selected granted written permission to conduct the study at their site. The two hospitals that were not able to grant permission in time were replaced with hospitals that were close by and had a similar staff-size.

Study enumerators spent 1-2 days at each study site where they administered a study questionnaire to the 40 pre-selected participants after obtaining a written informed consent. The study questionnaire elicited information on demographics, symptoms, testing, contacts and travel history. Enumerators also obtained 4-5 ml of blood samples from each participant. Blood samples were then transported to the Biochemistry Labs at Tribhuvan University Teaching Hospital (TUTH), and tested for antibodies against SARS-CoV-2.
Antibody test and validation

A Chemiluminescence Immunoassay (CLIA) based antibody test from Ortho Clinical Diagnostics (Vitros CoV2G [IgG]) was used to assess for the presence of antibodies to establish past exposure to SARS-CoV-2 among study individuals. The test detects IgG antibodies against S1 spike protein of SARS-CoV-2. We performed a local validation of the antibody test with 77 positive controls and 65 negative controls. Positive controls were serum samples from individuals with a positive polymerase chain reaction (PCR) test more than three weeks prior and negative controls were frozen serum samples that were obtained before December 2019. Manufacturer recommended test sample:calibrator optical signal ratio of 1.0 or more was used to identify positive results. We calculated the sensitivity and specificity of the serology test by fitting a Beta-Binomial Bayesian model using data from our validation study. We also used validation data submitted by the test manufacturer to regulatory agencies to generate strongly informative priors for the Beta-Binomial model. Based on the data, the sensitivity and specificity of the test in our context were 89.3% (95% Credible Interval (CrI), 85.8–93.0) and 99.2% (95% CrI 98.1%–99.8%) respectively.

Statistical analysis

The statistical analysis aimed to account for test accuracy in calculating the population based seroprevalence, to account for the effects of hospital based clustering, and to make the study findings representative of the study population. To do the first, we modeled the serology test result as a Bernoulli process. We then used Bayes’ rule to account for the test inaccuracy by populating it with measures of test sensitivity and specificity. To account for clustering at the hospital, we extended this model to a hierarchical Bayesian logistic regression model with partial pooling. To ensure representativeness of the study, we further extended this model into a multilevel (or hierarchical) regression model with post-stratification (MRP) by including age, gender, and occupation as predictors. We then weighted estimates of seroprevalence by the proportionate weight of these predictors to calculate the final
seroprevalence among health workers in Kathmandu. The statistical framework for our analysis is represented as below:(2,11–15)

\[ x_i \sim \text{Bernoulli}(p_i \ast \text{sens} + (1 - p_i) \ast (1 - \text{spec})) \]

\[ p_i = \text{logit}^{-1}(\alpha + \alpha_h \ast \sigma + x_i \beta) \]

\[ p \sim \text{Normal}(0.2, 1) \]

\[ \alpha \sim \text{Normal}(0, 1) \]

\[ \alpha_h \sim \text{Normal}(0, 1) \]

\[ \text{sens} \sim \text{Beta}(71, 9) \]

\[ \text{spec} \sim \text{Beta}(440, 2) \]

\[ x^+ \sim \text{Binomial}(n^+, \text{sens}) \]

\[ x^- \sim \text{Binomial}(n^-, \text{spec}) \]

\[ p(y_{pred} | y) \sim \int_0^1 p(y_{pred} | \theta) \ast p(\theta | y) \, d(\theta) \]

Relative Risk (RR) for group \( g \):

\[ (RR_g) = \frac{p_g}{p} = \frac{\text{logit}^{-1}(\alpha + \beta_0 + \beta_g + \alpha_h \ast \sigma)/\text{logit}^{-1}(\alpha + \beta_0 + \alpha_h \ast \sigma)}{1} \]

Here, \( x_i \) is the result of the sero-survey for the \( i^{\text{th}} \) individuals, \( p_i \) is the true underlying probability of a positive test for the \( i^{\text{th}} \) individual, \( \text{sens} \) is the test sensitivity, \( \text{spec} \) is the test specificity, \( \alpha \) is the fixed intercept term, \( \sigma \) is the standard deviation for the hospital random effect and, \( \alpha_h \) is the extent of deviation of the random effect in terms of \( \sigma \). \( x_i \) is a vector of predictor variables (age group, gender and occupation) and \( \beta \) is a vector of their respective coefficients. \( \theta \) represents the fitted parameters, \( y_{pred} \) is the new predicted data. Based on previous findings, we assumed a weakly informative normally distributed prior for the overall seroprevalence with a
mean of 0.2 and a standard deviation of 1. We created 40 strata (4 age categories x 2 gender categories x 5 occupation categories), and calculated seroprevalence for each of these 40 strata, which we then multiplied by their respective population weights to obtain the final seroprevalence.

We implemented this probabilistic model in the Stan programming language and interfaced it in R (version 4.0.3), via the Rstan package. Stan samples the posterior parameter space using Hamiltonian Monte Carlo (HMC) No U-Turn Sampler (NUTS). We ran 4 chains with 5000 iterations per chain and discarded the first 1000, resulting in 16,000 sampling iterations. To assess for model convergence, we used the R hat statistic, the number of effective samples, the energy parameter and visual measures. Visual model diagnostics are given in the supplementary appendix.

Overall and hospital-wise seroprevalence among health workers is reported as the mean and the 95% Credible Interval (CI) of the conditional probability of seropositivity given the data. Effect sizes are reported in terms of odds or relative risks. While calculating relative risk, the largest groups (age group 18–29, female gender, and nurses) were considered the respective reference groups.

**Results**

821 participants from 20 clusters (hospitals) participated in the survey. For each of the 20 clusters, we included the first 40 participants based on their order in the randomized sampling list, resulting in 800 records in our analytic sample. Of these 800 individuals, 522 (65.2%) were female, 372 (46%) were between ages 18-29, and 7 (0.9%) were 60 or above. 287 (36%) of the participants were nurses, 172 (22%) were administrative staff, 147 (18%) were doctors, 56 (7%) were laboratory and pharmacy staff while 137 (17%) were other staff with clinical or bedside roles. 529 (66%) were married, 485 (61%) had a bachelor’s degree or higher while 6% had no formal education (Table 1).

In comparison, based on records at the health ministry, 63% of health workers in the Kathmandu valley are females, 42% are between ages 18-29, 3% are above age 60, 30% are nurses, 36% are administrative staff and 20% are doctors. Additional details are available in the supplementary appendix.
Table 1: Demographic characteristics of SEVID-KaV participants based on SARS-CoV-2 antibody status

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Overall (N = 800)</th>
<th>Negative (N = 479)</th>
<th>Positive (N = 321)</th>
<th>p-value</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Age Group</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>18-29</td>
<td>372</td>
<td>214 (58%)</td>
<td>158 (42%)</td>
<td>0.2</td>
</tr>
<tr>
<td>30-49</td>
<td>355</td>
<td>213 (60%)</td>
<td>142 (40%)</td>
<td></td>
</tr>
<tr>
<td>50-59</td>
<td>66</td>
<td>47 (71%)</td>
<td>19 (29%)</td>
<td></td>
</tr>
<tr>
<td>&gt;= 60</td>
<td>7</td>
<td>5 (71%)</td>
<td>2 (29%)</td>
<td></td>
</tr>
<tr>
<td><strong>Gender</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Female</td>
<td>522</td>
<td>305 (58%)</td>
<td>217 (42%)</td>
<td>0.3</td>
</tr>
<tr>
<td>Male</td>
<td>278</td>
<td>174 (63%)</td>
<td>104 (37%)</td>
<td></td>
</tr>
<tr>
<td><strong>Occupation</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Nurse</td>
<td>288</td>
<td>167 (58%)</td>
<td>121 (42%)</td>
<td>0.8</td>
</tr>
<tr>
<td>Doctor</td>
<td>147</td>
<td>90 (61%)</td>
<td>57 (39%)</td>
<td></td>
</tr>
<tr>
<td>Other bedside/patient-care role (e.g. patient transport)</td>
<td>137</td>
<td>79 (58%)</td>
<td>58 (42%)</td>
<td></td>
</tr>
<tr>
<td>Laboratory/Pharmacy</td>
<td>56</td>
<td>34 (61%)</td>
<td>22 (39%)</td>
<td></td>
</tr>
<tr>
<td>Administration (including security)</td>
<td>172</td>
<td>109 (63%)</td>
<td>63 (37%)</td>
<td></td>
</tr>
<tr>
<td><strong>Marital Status</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Married</td>
<td>529</td>
<td>319 (60%)</td>
<td>210 (40%)</td>
<td>0.2</td>
</tr>
<tr>
<td>Unmarried</td>
<td>266</td>
<td>155 (58%)</td>
<td>111 (42%)</td>
<td></td>
</tr>
<tr>
<td>Divorced/Separated/Widowed</td>
<td>5</td>
<td>5 (100%)</td>
<td>0 (0%)</td>
<td></td>
</tr>
<tr>
<td><strong>Education</strong></td>
<td></td>
<td></td>
<td></td>
<td>0.074</td>
</tr>
<tr>
<td>Illiterate</td>
<td>25</td>
<td>9 (36%)</td>
<td>16 (64%)</td>
<td></td>
</tr>
<tr>
<td>Literate but no formal education</td>
<td>23</td>
<td>11 (48%)</td>
<td>12 (52%)</td>
<td></td>
</tr>
<tr>
<td>Primary education (Grade 5 or below)</td>
<td>26</td>
<td>14 (54%)</td>
<td>12 (46%)</td>
<td></td>
</tr>
<tr>
<td>Secondary education (Grade 6 to 12)</td>
<td>241</td>
<td>145 (60%)</td>
<td>96 (40%)</td>
<td></td>
</tr>
<tr>
<td>Bachelor degree or higher</td>
<td>485</td>
<td>300 (62%)</td>
<td>185 (38%)</td>
<td></td>
</tr>
<tr>
<td><strong>Monthly Income</strong></td>
<td></td>
<td></td>
<td></td>
<td>0.6</td>
</tr>
<tr>
<td>Up to Rs 20,000</td>
<td>131</td>
<td>74 (56%)</td>
<td>57 (44%)</td>
<td></td>
</tr>
</tbody>
</table>
Table 2 presents the frequency of common COVID-19 symptoms since January 2020 among seropositive and seronegative individuals. The symptoms with the unadjusted odds of seropositivity greater than one (i.e. more specific) were a loss of smell (odds of 4.64), fever (3), shortness of breath 2.35), muscle ache, diarrhea, rash and joint pain. The unadjusted odds of seropositivity among individuals with cough were 0.95 and headache were 0.94.

Figure 1: Unadjusted seroprevalence of COVID-19 among health workers in the Kathmandu valley

Note: Dotted line represents the mean unadjusted overall seroprevalence. Cyan bars represent unadjusted seroprevalence for individual hospitals.
Only about 70% of seropositive individuals in our study had one or more of the specific symptoms. Even when cough (with odds of seropositivity less than 1), was included among the symptoms, only 75% of the seropositive individuals had at least one symptom. Although 597 of the 800 participants had at least one PCR test, 181 participants (i.e. 23%) had a positive PCR test in the past. Odds of seroconversion among health workers who had a positive test in the past were 3.02, while they were 0.34 among individuals who had a negative PCR. Among health workers who did not have a PCR test in the past, the odds of sero-conversion were 0.60 (corresponding to a probability of 37%).

### Table 2: COVID-19 related symptoms in SEVID-KaV study participants

<table>
<thead>
<tr>
<th>Symptom</th>
<th>Overall (N = 800)</th>
<th>Negative (N = 479)</th>
<th>Positive (N = 321)</th>
<th>p-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Fever</td>
<td>172</td>
<td>43 (25%)</td>
<td>129 (75%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Shortness of Breath</td>
<td>94</td>
<td>28 (30%)</td>
<td>66 (70%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Cough</td>
<td>277</td>
<td>140 (51%)</td>
<td>137 (49%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Sputum Production</td>
<td>77</td>
<td>42 (55%)</td>
<td>35 (45%)</td>
<td>0.3</td>
</tr>
<tr>
<td>Loss of Smell</td>
<td>141</td>
<td>25 (18%)</td>
<td>116 (82%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Headache</td>
<td>372</td>
<td>192 (52%)</td>
<td>180 (48%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Myalgia</td>
<td>257</td>
<td>104 (40%)</td>
<td>153 (60%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Diarrhoea</td>
<td>79</td>
<td>33 (42%)</td>
<td>46 (58%)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Rash</td>
<td>25</td>
<td>9 (36%)</td>
<td>16 (64%)</td>
<td>0.013</td>
</tr>
<tr>
<td>Joint Pain</td>
<td>115</td>
<td>49 (43%)</td>
<td>66 (57%)</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

n (%) - Pearson’s Chi-squared test

Percentages are in terms of the row total. Columns add up to more than the column total because many individuals reported more than one symptom. Any report of symptoms since January 2020, when the pandemic started, is recorded as a positive.

321 of the 800 health workers included in the analysis tested positive for COVID-19 antibodies. A pooled (non-hierarchical) model of the overall seroprevalence without adjusting for test accuracy (sensitivity and specificity)
resulted in an unadjusted seroprevalence of 40.2% (95% CrI 36.8–43.6%). A model with no pooling between hospitals, unadjusted for test accuracy resulted in hospital-wise seroprevalence that ranged between 28.61% (95% CrI 16.17%–43.08%) and 52.15% (95% CrI 37.40%–67.01%) (Figure 1). Adjusted for test accuracy, the seroprevalence estimate from the unpooled model ranged between 33.6% (95% CrI 18.5%–51.5%) to 62.2% (95% CrI 44.0%–80.7%). These results are presented in Table 2.

Table 3: Relative risk of COVID-19 seropositivity based on age, gender and health worker occupation

<table>
<thead>
<tr>
<th>Age Group</th>
<th>Mean Seroprevalence %</th>
<th>Relative Risk (95% Credible Interval)</th>
</tr>
</thead>
<tbody>
<tr>
<td>18-29</td>
<td>41.7</td>
<td>1</td>
</tr>
<tr>
<td>30-49</td>
<td>40.5</td>
<td>0.97 (0.81-1.2)</td>
</tr>
<tr>
<td>50-59</td>
<td>30.7</td>
<td>0.74 (0.48-1)</td>
</tr>
<tr>
<td>60&lt;=</td>
<td>36.6</td>
<td>0.88 (0.31-1.6)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Gender</th>
<th>Mean Seroprevalence %</th>
<th>Relative Risk (95% Credible Interval)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female</td>
<td>42.0</td>
<td>1</td>
</tr>
<tr>
<td>Male</td>
<td>39.0</td>
<td>0.94 (0.75-1.1)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Occupation</th>
<th>Mean Seroprevalence %</th>
<th>Relative Risk (95% Credible Interval)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Nurse</td>
<td>41.7</td>
<td>1</td>
</tr>
<tr>
<td>Administration</td>
<td>39.0</td>
<td>0.94 (0.72-1.2)</td>
</tr>
<tr>
<td>Bedside support</td>
<td>44.3</td>
<td>1.1 (0.83-1.3)</td>
</tr>
<tr>
<td>Doctor</td>
<td>40.8</td>
<td>0.98 (0.74-1.3)</td>
</tr>
<tr>
<td>Laboratory/Pharmacy Personnel</td>
<td>41.8</td>
<td>0.99 (0.69-1.3)</td>
</tr>
</tbody>
</table>

*Reference group. Administration includes personnel that do not have a direct patient care responsibility including security personnel. Beside support refers to nurses aides and patient transporters.
Hospital-wise seroprevalence calculated from the fitted parameters generated from the final multilevel model with post-stratification was between 38.1% (95% CrI 30.7%–44.1%) and 40.5% (95% CrI 34.7%–47.0%) (Figure 2). Overall seroprevalence based on the final multilevel model with post-stratification, adjusted for test sensitivity and specificity was 38.99% (95% CrI 29.08%–43.91%). When weighted based on the age group, gender and occupation of health workers in the Kathmandu valley, the seroprevalence was 38.17% (95% CrI 29.26%–47.82%). Relative risk of seropositivity was the greatest among 18–29 year olds, females and bedside care providers, however many of these differences did not achieve statistical significance at the 95% credible interval (Table 3).

Figure 2: Adjusted hospital-wise seroprevalence of COVID-19 among health workers in the Kathmandu valley

Note: Red dots represent the mean seroprevalence for each hospital. Cyan error bars represent the 95% central predictive interval of the seroprevalence for each hospital. The dotted line represents the overall adjusted seroprevalence.
Discussion

Our analysis of the prevalence of antibodies against SARS-CoV-2 among hospital-based health workers in the Kathmandu valley, Nepal reveals a substantial exposure to the infection by the winter of 2020. By mid-December 2020, about 40% of the health workers had already developed antibodies against SARS-CoV-2, well before vaccination against COVID-19 had begun in Nepal. This means that these health workers had already been exposed to SARS-CoV-2 by the end of November, 2020 and subsequently developed antibodies against it. In addition, our findings show that seroprevalence is marginally higher (trending toward statistical significance) among health workers who have a direct patient care role (nursing and bedside clinical support roles). COVID-19 seroprevalence also appears to be inversely correlated—albeit weakly—with socioeconomic indicators (educational attainment and financial status). Although a large proportion of seropositive health workers experienced at least one symptom that was consistent with COVID-19, the most specific symptoms were a loss of sense of taste or smell, fever and shortness of breath. As expected, a positive PCR test in the past increased the odds of seropositivity substantially.

Because of the nature of their work, health care workers were thought to have a potentially higher risk of exposure to SARS-CoV-2. There are reasons for such expectations. Early in the epidemic, several countries, including Nepal, faced a shortage of personal protective equipment. This might have led to greater workplace exposure to SARS-CoV-2 among health workers. In addition, early in the epidemic there was also an inadequate understanding of the risk and mode of transmission (for e.g. transmission from asymptomatic individuals and airborne transmission or higher risk of transmission in unventilated closed spaces), this could have resulted in high risk of infection among health and other frontline workers.

Although the infection burden among health workers as shown by our study is substantial, it is not clear the extent to which this burden is different from the disease burden at the level of the community. Several studies have shown that seroprevalence among health workers often tracks seroprevalence in the community. And contrary to earlier expectations, at least a few studies have shown that in hospitals where adequate infection control practices are in place, health workers have a low risk of contracting the infection in the workplace. Therefore, it is unclear whether the seroprevalence seen in our study, although substantial, is the consequence of risk of exposure at
the workplace or the high community burden of the infection in Kathmandu. Across densely populated urban
communities of South Asia, there appears to have been a significant spread of COVID-19 within the first year of the
pandemic. An as yet unpublished estimate indicates that at least 17% of the overall population in and around
Kathmandu may have already been infected by September 2020. (24) In pockets of urban India, where COVID-19
related epidemic dynamics are similar in many ways to Kathmandu, this proportion was found to be even higher.
(25–27)

The 38% seroprevalence among the valley’s 25,000 or so health workers indicates that about 10,000 had contracted
the infection by the end of November, however official reports indicate that until then, only about 2500 of them had
been diagnosed. (28) This is even while health workers have comparatively better access to testing—many facilities
in Kathmandu routinely test their staff on a periodic basis. In fact, about 80% of the health workers in our sample
had already had at least one PCR test as part of routine surveillance. The four fold gap between cases and infections
is partly explained by the fact many individuals who contracted the infection appear to have developed no symptoms
at all, or experienced mild symptoms for which they did not seek testing. In the general population, the gap between
cases and infections could be even higher as they have poorer access to testing services, or may not seek care in the
first place.

Our study has several strengths. First, the study was designed to be representative of all the hospital based health
workers in the Kathmandu valley. As our comparison with the overall health worker population of Kathmandu
shows, our sampled population appears to well represent Kathmandu’s health workers based on their age group,
gender and occupational group. In addition, our study post-stratified and weighted seroprevalence based on these
demographic variables to make the findings representative of the study population. The fact that post-stratification
resulted in less than a percentage difference between the unweighted and weighted seroprevalence means that our
study sample was remarkably representative of the overall health worker population. Our methodology allowed for
an easy correction of test accuracy. The study sampled more than 3% of the study population. The comparatively
large sample size, combined with the hierarchical statistical framework allowed for more accurate estimates of
seroprevalence. The longitudinal cohort design of this study means that it will allow us to study not only the
temporal variation of seroprevalence but also to study antibody decay in the future.
A major limitation of this study is our measure of test accuracy. Our unadjusted measure for sensitivity was 82% while it was 90% based on the manufacturer’s data alone. It is possible that some of the PCR test results that we used to identify positive controls were falsely positive, especially among individuals who did not have a clinical diagnosis of COVID-19. To account for this limitation, we derived strongly informative Bayesian priors from the manufacturer’s data and allowed them to influence our final calculation of test accuracy. Because of this, our measure of the test accuracy in the final model—with an 89% sensitivity and greater than 99% specificity closely matches manufacturer’s data. The added benefit of our validation data may be that it may reflect upon local testing conditions. Surveys like these might be biased because individuals who agree to participate in such surveys may have a greater tendency to seek care, or may be at a higher risk. In our survey, 40% of those invited agreed to participate and an overwhelming majority of those who declined cited scheduling conflict as the reason for not participating. In addition, unadjusted seroprevalence among those who had not had a PCR test in the past (correlating to health care seeking behaviour and risk), was 37%, meaning that our survey sample was fairly well balanced.

**Conclusions**

A significant proportion of health workers in Kathmandu appear to have been infected with COVID-19 by the end of 2020. Although it is not entirely clear to what extent health workers were infected at the workplace, these seroprevalence figures still warrant a reassessment of infection control practices at Kathmandu’s hospitals. If health worker and community seroprevalence are correlated, these estimates may also be indicative of the seroprevalence in the overall community. Since seroprevalence studies are easier to conduct among health workers than in the community, they could be used to inform the serostatus of the overall community and predict future epidemic dynamics and disease spread. Future waves of this study will be useful in assessing the progress of the epidemic over time, the temporal variation of antibodies, and now that a sizable proportion of health workers have been vaccinated, they could also generate insight on the real world evidence of vaccine efficacy.


Additional Information

Authors’ Contributions

- Designed the study: KRP, AB, SP, AS, PA, PG
- Conducted the study: KRP, AB, SP, RB, JP, AS, PA, DA, PJP, GSS, KP, NT, ST
- Laboratory investigations: AB
- Software and data analysis: KRP
- First draft: KRP
- Review and final draft: KRP, AB, SP, RB, JP, AS, PA, DP, PJP, GSS, KP, NT, ST, PG

Competing interests

None of the authors have any competing interests to declare.

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Supplementary Files

This is a list of supplementary files associated with this preprint. Click to download.

- SEVIDKaVW1SuppAppendix.pdf