Preexisting autoimmune disease is a risk factor for immune-related adverse events: a meta-analysis

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Research Article

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Abstract

Purpose Patients with preexisting autoimmune disease (PAD) are often excluded from clinical trials assessing immune checkpoint inhibitors (ICIs). Therefore, the safety of ICI therapy in patients with PAD remains unclear. Herein, we evaluated the incidence of immune-related adverse events (irAEs) in patients with PAD when compared with non-PAD patients.

Methods We searched MEDLINE/PubMed, Web of Science, and Google Scholar for eligible studies from inception to January 2021. Observational studies reporting the incidence of irAEs in patients with and without PAD were included. We then performed a meta-analysis of eligible studies using forest plots. The primary endpoint of this study was the incidence rate of irAEs between patients with and without PAD.

Results We identified three prospective and three retrospective studies involving 206 patients with PAD and 3078 patients without PAD. In the meta-analysis, 128 patients with PAD (62.1%) experienced irAEs, which occurred in 51.9% of non-PAD patients, resulting in an odds ratio (OR) of 2.14 (95% confidence interval [CI] 1.58–2.89). In the subgroup analysis, the incidence of irAEs was significantly higher in patients with PAD (OR = 2.19, 95% CI [1.55–3.08]). Furthermore, no significant heterogeneity or publication bias was detected, indicating that our meta-analysis could be generalized to clinical settings.

Conclusion This meta-analysis demonstrated that PAD was a risk factor for irAE incidence. These results suggest that monitoring the occurrence of irAEs in patients with PAD is required to manage irAEs appropriately.

1. Introduction

Immune checkpoint inhibitors (ICIs) have remarkable anticancer activity for treating several types of cancer [1]. ICIs block the immune system-mediated response toward tumors by interacting between cytotoxic T-lymphocyte-associated-4 (CTLA-4) and CD80/CD86, as well as between programmed cell death protein 1 (PD-1) and programmed cell death protein ligand 1 (PD-L1) [1]. To date, anti-CTLA-4 antibody (ipilimumab), anti-PD-1 antibodies (nivolumab and pembrolizumab), and anti-PD-L1 antibodies (atezolizumab, durvalumab, and avelumab) have been approved, and their clinical applications have dramatically expanded.

Despite the significant efficacy of ICIs, adverse events associated with ICIs, called immune-related adverse events (irAEs), have been reported [2]. IrAEs are associated with the mechanism of ICI action and can involve multiple organs, including the skin, endocrine system, and gastrointestinal tract [3]. Typically, irAEs are mild, and ICI therapy can be continued in most cases. However, severe irAEs result in treatment discontinuation and reduce the quality of life. Therefore, adequate management of irAEs is required to maximize the benefits of ICI therapy [3].

Although the pathophysiology of irAEs is not well understood, the underlying mechanisms could involve upregulation of the immune system, which, in turn, causes inflammation and autoimmunity, similar to autoimmune diseases such as rheumatoid arthritis, autoimmune thyroid disease, and inflammatory bowel disease [4]. Accordingly, several clinical trials have excluded these populations from study participation. Hence, the efficacy and safety of ICIs in these patients remains unknown. Recently, a meta-analysis of observational studies assessing preexisting autoimmune disease (PAD) populations has reported that ICI treatments are effective and irAEs are often manageable [5]. However, the study included PAD single-arm studies and did not compare the incidence of irAEs between patients with and without PAD. Furthermore, only a few comparative studies have been reported, and the results obtained so far remain controversial.
Based on this background, in the present study, we aimed to reveal whether PAD induces irAEs, by performing a meta-analysis of observational studies that reported the incidence of irAEs, with and without PAD incidence.

2. Methods

2.1. Study design

We performed a meta-analysis according to the Preferred Reporting Items for Systematic reviews and Meta-Analyses (PRISMA) [6]. The primary outcome of this meta-analysis was the incidence of irAEs between patients with and without PAD.

2.2. Data sources and searches

We developed a search strategy to identify related records, with no language restriction, for MEDLINE/PubMed, Web of Science, and Google Scholar, from database inception to January 31, 2021. The search terms included neoplasm, immune checkpoint inhibitors, and autoimmune disease. The titles and abstracts of the searched articles were screened, and the full text was assessed for eligible studies. Finally, we selected studies that included extractable data for the incidence rate of irAEs in patients with or without PAD.

2.3. Eligibility criteria

Studies were considered eligible if they were observational studies and reported the number of patients receiving ICIs and the incidence of irAEs with and without PAD. We excluded case reports/case series, review articles, or studies including the PAD-alone cohort.

2.4. Data extraction

Data extraction was independently conducted by two authors (AY and MK). We extracted the following data: author, publication year, study design, study period, cancer type, therapeutic agent, characteristics of PAD, and incidence rate of irAEs.

2.5. Data synthesis and statistical analysis

Data were analyzed using Review Manager 5.4 software (The Cochrane Collaboration, Copenhagen, Denmark) and R (R Foundation for Statistical Computing, Vienna, Austria). Meta-analysis was performed using the fixed effects model and the Mantel-Haenszel (M-H) method. We calculated the odds ratio (OR) and 95% confidence interval (95%CI). Statistical significance was defined as a Z index p-value < 0.05. The heterogeneity among studies was assessed using the Cochran Q test (Chi²) and the I² statistics: Chi² p-value < 0.1 correlated high heterogeneity; I² < 25%, 25–50%, and > 50% indicated low, moderate, and high heterogeneity, respectively. To explore the possibility of publication bias, we generated funnel plots (visual inspection) and performed Begg’s test (rank correlation) and Egger’s test (linear regression); an asymmetric funnel plot and a p-value < 0.1 implied potential publication bias [7, 8].

3. Results

3.1. Study selection and characteristics

A systematic search of articles identified 600, 552, and 327 reports in MEDLINE/PubMed, Web of Science, and Google Scholar, respectively. After screening these reports, we assessed 32 full-text articles and included 6 articles in this meta-analysis (Fig. 1). In total, 206 patients with PAD and 3078 patients without PAD were evaluated. The study characteristics are summarized in Table 1. Among the identified studies, three were retrospective studies [9–11], and the others were prospective studies [12–14]. The incidence of irAEs was the primary endpoint in two studies [10, 11]
and secondary in the others [9, 12–14]. In four studies, the patients were mainly treated with an anti-PD-1 antibody (nivolumab, pembrolizumab) [9, 10, 12, 14], and the two remaining studies were based on mixed ICIs (ipilimumab and anti-PD-1 antibody) [11] and atezolizumab [13]. The types of PAD covered a broad range, including rheumatoid arthritis, endocrine disorders, dermatologic disorder, and gastrointestinal disorders. Five studies reported irAEs of any grade [9–11, 13, 14], and one study reported grade 2 or more irAEs [12].

3.2. Quantitative synthesis and meta-analysis

According to the forest plot of the six studies, the occurrence of all reported irAEs was 62.1% in patients with PAD and 51.9% in non-PAD patients. In addition, the incidence of irAEs was significantly higher in patients with PAD than in non-PAD patients (OR = 2.14, 95% CI [1.58–2.89], Z = 4.94, p < 0.05) (Fig. 2). The Chi² p-value was 0.45 and I² was 0%, indicating that the heterogeneity was low in this meta-analysis.

3.3. Risk of publication bias

The Begg's test (z = 0.19, p = 0.85) and Egger's test (t = -0.76, p = 0.49) suggested no publication bias. Although the number of studies was extremely limited to apply funnel plots and data were lacking in the large SE area, visual inspection of the funnel plot might exhibit symmetry (Fig. 3).

3.4. Subgroup analysis

Next, we conducted a subgroup analysis of irAEs of any grade for the early detection of irAEs. We eliminated one study that reported grade 2 or more irAEs to unify the outcome and exclude the effects of grade 1 irAEs on our main meta-analysis (Fig. 4). The forest plot indicated a similar result as described in all reported irAEs (67.0% in PAD vs. 54.8% in non-PAD; OR = 2.19, 95% CI [1.55–3.08], Z = 4.47, p < 0.05). Heterogeneity was also low (Chi² = 4.61, p = 0.33, I² = 13%). Based on these results, we suggest that PAD increases the risk of irAEs of any grade.

4. Discussion

In the present study, we conducted a meta-analysis of six reports to elucidate the influence of ICI therapy on irAEs in patients with PAD. We observed that PAD was significantly associated with a higher risk of irAE incidence (OR = 2.14). The incidence of irAEs in the PAD group was 62.1%, which was similar to a previous pooled analysis (60%) that included PAD single-arm observational studies [5]. Therefore, we believe that the population of this study was relevant in evaluating the risk of irAEs. Furthermore, the results of our subgroup analysis were similar to those of the main analysis. The heterogeneity among studies was low, and the integrated effects were significant, suggesting that the results of our meta-analysis could be generalized. Our statistical test (Begg's test and Egger's test) revealed that publication bias was low. In the funnel plot, it was difficult to assess symmetry owing to a lack of data on the base side; therefore, additional studies are needed.

In a previous report, Bender et al. have researched hospitalization rates, which implied clinically severe cases, in patients with melanoma receiving PD-1 inhibitors and showed that PAD significantly increased the hospitalization rates when compared with non-PAD cases (24.1% vs. 5.8%, p < 0.0001) [15]. In another study, PAD was reportedly associated with hospitalization following irAE diagnosis (HR 1.81, 95%CI [1.21–2.71], p < 0.05) [16]. These reports support our results and suggest that irAE management in patients with PAD is more challenging than that in non-PAD patients, necessitating additional monitoring of symptoms and statements for appropriate ICI therapy. In the present study, the association between PAD and severe irAEs was not evaluated. Thus, further studies, such as a meta-analysis of hospitalization in patients with PAD during ICI treatment, are required.
In clinical settings, an autoimmune disease can develop owing to an increase in autoantibodies such as rheumatoid factor in rheumatoid arthritis, antithyroglobulin antibody, thyroid-stimulating hormone (TSH) receptor antibody in autoimmune thyroid disease, and antinuclear antibody (ANA) in connective tissue disease. Hence, it is considered that preexisting autoantibodies are also associated with irAE incidence and may be valuable clinical biomarkers. Indeed, a study was conducted to assess whether preexisting autoantibodies were associated with the safety and efficacy of ICI treatment in patients with non-small cell lung cancer and revealed that these preexisting antibodies were associated with irAE incidence, and ICI efficacy was higher in patients with any of the preexisting antibodies [17]. The study also supports our results that PAD increases the risk of irAEs, although autoantibodies do not consistently result in an autoimmune disease. Conversely, ANA was not associated with any grade irAEs, although only a small number of patients were evaluated [18]. The efficacy tended to be superior in patients without ANA, which was not statistically different [18]. Further research is required to elucidate the association between PAD, preexisting autoantibodies, and irAE incidence.

In the present study, Cortellini et al. reported that the incidence rate of irAEs of any grade was significantly higher in patients with inactive and active PAD than in those without PAD (64.3%, 73.3%, and 39.9%, \( p = 0.0001 \) and \( p = 0.0402 \), respectively) [10]. This rate was similar to that observed in the systematic review of patients with inactive and active PAD (67% and 75%, respectively) [19]. Based on these findings, it can be suggested that the incidence of irAEs is related to PAD, regardless of the baseline autoimmune disease activity.

In general, autoimmune diseases are caused by environmental and genetic factors. The accumulation of these triggers induces the breakdown of immune tolerance and the production of autoreactive B cells, T cells, and autoantibodies, which lead to excessive activation of both innate and adaptive immune cells. Finally, the aberrant immune system induces chronic inflammation and organ-specific or systemic damage [20, 21]. Additionally, autoimmune diseases cause various systemic symptoms. For example, inflammatory bowel disease presents extraintestinal manifestations such as in the joints, skin, or eyes [22], and rheumatoid arthritis is a risk factor for interstitial lung disease [23], indicating that the immune reactions occur in various tissues in autoimmune diseases, like irAEs. Consequently, we propose a potential mechanism of irAE development in patients with PAD. First, patients with PAD present systemic immune state instability, although clinical symptoms may be absent. The administration of ICIs blocks the immune checkpoint pathways and enhances T-cell effector function. Finally, T-cell activation can easily disrupt the balance of the immune system owing to the fragility of the host immune state, inducing flares of autoimmune disease and systemic irAEs. However, the molecular mechanisms involved, such as regulatory T cells, interleukin (IL), and tumor necrosis factor-alpha (TNF-\( \alpha \)), remain elusive. Further basic research is required to reveal detailed underlying mechanisms.

To avoid irAEs in patients with PAD, immunosuppressive therapy during the early stages may be effective. In the previously discussed systematic review, patients receiving immunosuppressive therapy at the initiation of ICI treatment demonstrated a marginally lower incidence rate of irAEs than those without treatment (67% vs. 74%) [19]. In contrast, baseline corticosteroid use (\( \geq 10 \text{ mg} \)) was associated with worse progression-free survival (hazard ratio, 1.31; 95% CI [1.03–1.67], \( p = 0.03 \)) and overall survival (hazard ratio, 1.66, 95% CI [1.28–2.16], \( p < 0.01 \)) [24]. Therefore, it is necessary to consider corticosteroid administration in patients with PAD depending on symptom severity.

This study has some limitations. First, the number of included reports was small, and four out of six studies used secondary outcome data. Therefore, the quality of evidence was low, and the evaluation of the funnel plot symmetry might have been insufficient, indicating a potential publication bias. Second, the treatment type (ICI monotherapy or combination with cytotoxic chemotherapy or molecular targeted therapy) was not described in some studies. Third,
patient characteristics (cancer type, treatment agent, and PAD type) were not unified, and baseline characteristics
were not balanced in the meta-analysis. In general, autoimmune diseases are common in women, which may affect
the incidence rate of irAEs, although baseline characteristics did not affect irAE incidence [25]. Additionally, we were
unable to determine the relationship between these factors and irAEs. Fourth, the severity of irAEs was not evaluated.
Although the frequency of irAEs and hospitalization rates were increased in patients with PAD, further investigations
on severe irAEs are required. Finally, the reported irAEs in patients with PAD might have included PAD flare and de
novo irAEs, which were not described in detail in the included studies. It is difficult to assess whether the PAD flare
could be attributed to ICI treatment. Thus, we might have overestimated irAEs in patients with PAD.

In conclusion, we performed a meta-analysis and identified that PAD is a risk factor for irAE incidence. We believe
that this study suggests the importance of carefully monitoring patients with PAD receiving ICIs for early detection
with appropriate countermeasures. Accordingly, the establishment of co-treatments and suitable strategies to
promptly collect information regarding patient conditions are required.

**Declarations**

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None.

**Conflicts of interest/Competing interests:** (include appropriate disclosures)

The authors have no conflicts of interest and competing interests to declare.

**Availability of data and material:** (data transparency)

The dataset used for this study are available from the corresponding author on reasonable request.

**Code availability:** (software application or custom code)

Not applicable.

**Authors’ contributions:** (include appropriate statements)

AY, YS, and MK designed the study. AY and MK conducted the study and analyzed the data. AY, YS, MK, KO, KN, AF,
YT, and MS contributed to the writing of the manuscript. All authors reviewed the results and approved the final
version of the manuscript.

**Ethics approval:** (include appropriate approvals or waivers)

Not applicable.

**Consent to participate:** (include appropriate statements)

Not applicable.

**Consent for publication:** (include appropriate statements)

Not applicable.
References


Table 1 Characteristics of the eligible studies
<table>
<thead>
<tr>
<th>Study</th>
<th>Study design and period</th>
<th>Cancer type</th>
<th>Therapeutic agent</th>
<th>Characteristics of PAD</th>
<th>irAEs grade</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>Nivolumab (n = 52)</td>
<td>Psoriasis (n = 1) Unknown (n = 1)</td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>Pembrolizumab (n = 1)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cortellini 2019 [10]</td>
<td>Retrospective Sep 2013 - May 2018</td>
<td>Advanced cancer (stage IV)</td>
<td>Anti PD-1 antibody</td>
<td>Thyroid disorders (n = 51) Dermatologic (n = 14)</td>
<td>Any grade irAEs</td>
</tr>
<tr>
<td></td>
<td></td>
<td>NSCLC (n = 492) Melanoma (n = 159)</td>
<td>Pembrolizumab (n = 182)</td>
<td>Rheumatologic (n = 10) Gastrointestinal/hepatic (n = 4)</td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td>RCC ( n = 94) Others (n = 6)</td>
<td>Nivolumab (n = 569)</td>
<td>Neurologic (n = 1) Nephroligic (n = 1) Multiple site (n = 4)</td>
<td></td>
</tr>
<tr>
<td>Danlos 2018 [12]</td>
<td>Prospective Jun 2014 - Dec 2016</td>
<td>Melanoma (n = 36) NSCLC (n = 6)</td>
<td>Pembrolizumab (n = 34) Nivolmab (n = 10) Avelumab (n = 3)</td>
<td>Rheumatic (n = 7) Dermatologic (n = 33) Endocrine (n = 9) Neurologic (n = 3) Haematologic (n = 1)</td>
<td>≥ Grade 2 irAEs</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Other cancers (n = 3) (Only PAD cohort was described)</td>
<td>Only PAD cohort was described</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Loriot 2020 [13]</td>
<td>Prospective Nov 2016 - Mar 2018</td>
<td>Locally advanced/metastatic urothelial carcinoma or nonurothelial carcinoma (n = 997)</td>
<td>Atezolizumab (n = 997)</td>
<td>Psoriasis (n = 15) Rheumatoid arthritis (n = 4) Hypothyroidism or Thyroiditis (n = 4) Ulcerative colitis (n = 2) Hashimoto’s disease (n = 2) Sarcoidosis (n = 2) Graves disease (n = 2) Others (n = 4)</td>
<td>Any grade irAEs</td>
</tr>
<tr>
<td>Schadendorf 2019 [14]</td>
<td>Prospective Aug 2017 - Jan 2019</td>
<td>stage III or IV melanoma (n = 1008)</td>
<td>Nivolumab (n = 1008)</td>
<td>Endocrine (n = 14) Gastrointestinal (n = 2) Hepatic (n = 1) Skin (n = 7) Other (n = 1)</td>
<td>Any grade irAEs</td>
</tr>
</tbody>
</table>

Abbreviations: R/R cHL, relapsed/refractory classical Hodgkin lymphoma; NSCLC, non-small cell lung cancer; RCC, renal cell carcinoma; PD-1, programmed cell death protein 1; PD-L1, programmed cell death protein ligand 1; PAD, pre-existing autoimmune disease.

**Figures**
Figure 1

Flow diagram of research strategy and study selection.

<table>
<thead>
<tr>
<th>Study or Subgroup</th>
<th>PAD Events</th>
<th>Total</th>
<th>non-PAD Events</th>
<th>Total</th>
<th>Weight</th>
<th>Odds Ratio M-H, Fixed, 95% CI</th>
<th>Odds Ratio M-H, Fixed, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bair 2019</td>
<td>3</td>
<td>4</td>
<td>25</td>
<td>48</td>
<td>1.6%</td>
<td>2.88 [0.26, 29.64]</td>
<td></td>
</tr>
<tr>
<td>Cartellini 2019</td>
<td>56</td>
<td>85</td>
<td>296</td>
<td>666</td>
<td>34.6%</td>
<td>2.90 [1.81, 4.67]</td>
<td></td>
</tr>
<tr>
<td>Darlots 2018</td>
<td>20</td>
<td>45</td>
<td>102</td>
<td>352</td>
<td>21.6%</td>
<td>1.86 [1.04, 3.69]</td>
<td></td>
</tr>
<tr>
<td>Kardio 2018</td>
<td>8</td>
<td>12</td>
<td>33</td>
<td>66</td>
<td>5.7%</td>
<td>2.68 [1.55, 7.23]</td>
<td></td>
</tr>
<tr>
<td>Lorenzo 2020</td>
<td>24</td>
<td>35</td>
<td>506</td>
<td>962</td>
<td>18.8%</td>
<td>1.87 [0.95, 4.08]</td>
<td></td>
</tr>
<tr>
<td>Schaedercort 2018</td>
<td>17</td>
<td>25</td>
<td>604</td>
<td>883</td>
<td>17.7%</td>
<td>1.02 [0.44, 2.39]</td>
<td></td>
</tr>
</tbody>
</table>

Total (95% CI): 206/3078 = 100.0% 2.14 [1.53, 2.89]

Total events: 128/1596

Heterogeneity: Chi² = 4.70, df = 5 (P = 0.45), I² = 0%

Test for overall effect: Z = 4.94 (P < 0.001; CI)

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Figure 2

Forest plot of meta-analysis of all reported immune-related adverse events (irAEs).
Figure 3
Funnel plot of eligible studies.

<table>
<thead>
<tr>
<th>Study or Subgroup</th>
<th>PAD Events</th>
<th>Total Events</th>
<th>non-PAD Events</th>
<th>Total Events</th>
<th>Weight</th>
<th>Odds Ratio M-H, Fixed, 95% CI</th>
<th>Odds Ratio M-H, Fixed, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bair 2019</td>
<td>3</td>
<td>4</td>
<td>25</td>
<td>48</td>
<td>2.0%</td>
<td>2.80 [0.26, 29.64]</td>
<td></td>
</tr>
<tr>
<td>Cornell 2019</td>
<td>56</td>
<td>85</td>
<td>286</td>
<td>666</td>
<td>44.1%</td>
<td>2.96 [1.81, 4.67]</td>
<td></td>
</tr>
<tr>
<td>Kanol 2018</td>
<td>8</td>
<td>12</td>
<td>33</td>
<td>66</td>
<td>7.3%</td>
<td>2.00 [0.56, 7.29]</td>
<td></td>
</tr>
<tr>
<td>Loriot 2020</td>
<td>24</td>
<td>35</td>
<td>506</td>
<td>862</td>
<td>24.0%</td>
<td>1.97 [0.96, 4.09]</td>
<td></td>
</tr>
<tr>
<td>Schadendorf 2013</td>
<td>17</td>
<td>25</td>
<td>694</td>
<td>983</td>
<td>22.6%</td>
<td>1.02 [0.44, 2.39]</td>
<td></td>
</tr>
<tr>
<td>Total (95% CI)</td>
<td>161</td>
<td>2726</td>
<td>100.0%</td>
<td></td>
<td></td>
<td>2.49 [1.55, 3.98]</td>
<td></td>
</tr>
<tr>
<td>Total events</td>
<td>108</td>
<td>1434</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Heterogeneity: Chi² = 4.61, df = 4 (P = 0.33), I² = 13%
Test for overall effect: Z = 4.47 (P < 0.00001)

Figure 4
Forest plot of meta-analysis of any grade irAEs (subgroup analysis).