

1 Spatio-temporal patterns of pneumonia in Bhutan: A Bayesian 2 analysis

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26 **Abstract**

27 Pneumonia is one of the top 10 diseases by morbidity in Bhutan. This study aimed to investigate
28 the spatial and temporal trends and risk factors of pneumonia in Bhutan. A multivariable Zero-
29 inflated Poisson regression using a Bayesian Markov chain Monte Carlo simulation was
30 undertaken to quantify associations of age, sex, rainfall, maximum temperature and relative
31 humidity with monthly pneumonia incidence and identify underlying spatial structure of the
32 data. Overall pneumonia incidence was 96.5 and 4.57 per 1,000 populations over nine years in
33 people aged <5 years and ≥ 5 years, respectively. Children <5 years or being a female are more
34 like to get pneumonia than ≥ 5 years and males. A 10mm increase in rainfall and 1°C increase
35 in maximum temperature was associated with a 7.2% (95% (credible interval [CrI] 0.7%,
36 14.0%) and 28.6% (95% CrI 27.2%, 30.1%) increase in pneumonia cases. A 1% increase in
37 relative humidity was associated with a decrease in the incidence of pneumonia by 8.6% (95%
38 CrI 7.5%, 9.7%). There was no evidence of spatial clustering after accounting for the
39 covariates. Seasonality and spatial heterogeneity can partly be explained by the association of
40 pneumonia risk to climatic factors including rainfall, maximum temperature and relative
41 humidity.

42 **Keywords:** Bhutan, pneumonia, Bayesian, spatial, temporal, risk factors, modelling

43 **Introduction**

44 Pneumonia is a major cause of morbidity and mortality worldwide¹. Each year, pneumonia
45 accounts for over 12 million hospital admissions and 1.3 million deaths in children aged less
46 than 5 years worldwide^{2,3}. In 2017, pneumonia was the fourth-leading cause of death and it is
47 estimated that it will be the third-leading cause of death by 2040⁴. The World Health
48 Organization (WHO) estimates that respiratory infections account for 6% of the total global
49 burden of disease. This accounts for a higher percentage compared to the burden of diarrheal
50 disease, cancer, human immunodeficiency virus (HIV) infection, ischemic heart disease or
51 malaria⁵.

52 Pneumonia is a potentially life-threatening illness with a particularly high burden in South Asia
53 and sub-Saharan Africa^{3,6,7}. It is not only a major cause of morbidity and mortality but is also
54 associated with a substantial economic burden on healthcare systems^{8,9} and household
55 income¹⁰. Pneumonia often has a complex aetiology involving multiple pathogens, including
56 many that are transmitted person-to-person. Past time-series analyses have identified various
57 pneumonia and influenza outcomes to be temporally seasonal, demonstrating highly consistent
58 peaks in winter months and troughs in summer months^{11,12}. Other studies have found that
59 pneumonia admissions were highly spatially clustered¹³, driven by contact with infected people
60 during indoor activities¹⁴.

61 Pneumonia continues to be an important communicable disease in Bhutan- locate in the Eastern
62 Himalayas¹⁵⁻¹⁷ (**Fig. 1**). In 2019, pneumonia was one of the top-ten ranked diseases in terms of
63 morbidity and accounted for 19% of the overall disease burden¹⁸. Every year the Bhutanese
64 government spends a huge amount on the treatment and management of pneumonia. In the
65 financial year 2017–2018, 7.1% of current health expenditure was spent on treating infectious
66 respiratory diseases^{19,20}. Despite the importance of pneumonia, and the infectious nature of the
67 disease, there have been no previous studies to understand the underlying ecological drivers of

68 pneumonia in the country^{21,22}. Understanding the spatial and temporal patterns of pneumonia
69 will be important for prevention and preparedness through more efficient targeting of scarce
70 health care resources. This study aims to investigate the trends of pneumonia, identify potential
71 high-risk geographical areas and quantify associations between disease risk and climatic risk
72 factors.

73 **Results**

74 **Descriptive analysis**

75 A total of 100,015 pneumonia cases were reported in the country during the study period (2010-
76 2018). This corresponded to 71,807 and 28,208 cases in people aged <5 years and ≥ 5 years,
77 with an incidence of 96.5 and 4.57 cases, respectively, per 1,000 people during the nine years
78 (**Table 1**). In both the age groups incidence decreased: from 119.28 and 47.73 cases per 1,000
79 population in 2010 to 54.73 and 3.19 cases per 1,000 in 2018 for the <5 years and ≥ 5 years age
80 groups, respectively (**Table 1**). The seasonal-trend decomposition of monthly pneumonia
81 cases based on locally (STL) is illustrated in Figure 2. The highest cases were reported in 2014
82 and pneumonia displayed a strong seasonal pattern. There were two peaks in May and
83 September of each year. The standard morbidity ratio (SMR) of pneumonia at sub-district level
84 varied from 0 to 13.02, with a Standard Deviation=1.45 (**Fig. 3**).

85 **Spatio-temporal model**

86 Model I, containing the unstructured random effects was better fitting than Model II and Model
87 III containing the spatially structured random effects with lower deviation information criterion
88 (DIC) (206,093). The incidence of pneumonia was 21.3% (95% credible interval [CrI] 21.0%,
89 21.6%) times higher in people aged <5 years as compared to ≥ 5 years. Females were 8% (95%
90 CrI 7.0%, 9.0%) more like to get pneumonia compared to males. Pneumonia decreased by 11%
91 (95% CrI 10%, 13%) during the study period. A 10mm increase in rainfall was associated with

92 a 7.2 % (95% CrI 0.7%, 14.0%) increase in incidence of pneumonia. Similarly, a maximum
93 temperature increase of a 1°C was associated with a 28.6% (95% CrI 27.2%, 30.1%) increase
94 in pneumonia cases. However, a 1% increase in relative humidity was associated with a
95 decrease in the incidence of pneumonia by 8.6% (95% CrI 7.5%, 9.7%) (**Table 2**).

96 There was no evidence of spatial clustering after accounting for the covariates (**Table 2** and
97 **Fig. 4**). There was >95% probability of a higher than the national average trend of pneumonia
98 in 56/205 sub-districts, whereas 67/205 sub-districts had >95% probability of a trend less than
99 the national average. There was no clear spatial pattern, with sub-districts showing higher and
100 lower average trends across all the 20 districts (**Fig. 5**).

101 102 **Discussion**

103 Pneumonia was spatially and temporally heterogeneous across sub-districts of Bhutan during
104 the study period. There was a decreasing trend, in addition to a strong seasonal pattern during
105 the study period. Pneumonia mainly affected children aged <5 years and females. Rainfall and
106 maximum temperature were associated with an increased incidence of pneumonia while
107 relative humidity was associated with a decrease incidence.

108 In addition to climatic factors, spatial heterogeneity could be due to differences in the socio-
109 demographic characteristics of sub-districts. The risk factors responsible for exacerbation and
110 spread of pneumonia in Bhutan were low birth weights, malnutrition, smoky and overcrowding
111 households, bottle-feeding of infants and poor personal and environmental hygiene^{23,24}. This
112 was evident from the two districts of Haa and Paro which has the lowest poverty²⁵, and also
113 reported lowest SMR of pneumonia. Similar to the decreasing trend in the incidence of global
114 childhood pneumonia²⁶, the national pneumonia trend decreased during the study period. This

115 could be attributed to a decrease in exposure to key risk factors including poor housing
116 conditions and overcrowding, incomplete immunisation and malnutrition²⁶.

117 Pneumonia is the single largest infectious cause of death in children worldwide. It accounts for
118 15% of all deaths of children <5 years²⁷. In this study, children <5 years were at a much higher
119 risk of pneumonia compared to those ≥ 5 years. Infants (aged between 0-11 months) was
120 reported to contribute up to 24.2% of cases in another study²⁸. The WHO and United Nations
121 Children's Fund (UNICEF) initiated a Global action plan for pneumonia and diarrhoea
122 (GAPPD) to accelerate pneumonia control in children²⁹. The GAPPD strategies include
123 promoting exclusive breastfeeding and adequate complementary feeding to protect children
124 from pneumonia; prevent pneumonia through vaccinations, hand washing with soap, reducing
125 household air pollution, HIV prevention and cotrimoxazole prophylaxis for HIV infected and
126 exposed children; and treating children with pneumonia with antibiotics and oxygen.
127 Strengthening GAPPD strategies should be considered in Bhutan, as is the case in other
128 countries in the South Asia region (Bangladesh and India). The introduction of pneumococcal
129 conjugate vaccines in Bhutan in 2019 is timely in prevention of pneumonia^{19,30}. Exclusive
130 breastfeeding rates from birth until six months in Bhutan varies from 35.9–51.0%^{31,32}.
131 Increasing exclusive breastfeeding rates are likely to reduce pneumonia associated morbidity³³.
132 Pneumonia was highly seasonal and was associated with climatic factors including
133 temperature, rainfall and relative humidity. The association of temperature with pneumonia has
134 been reported in other studies^{34,35}. A plausible explanation is the association of higher
135 temperature with air pollution³⁶ which in itself is known risk factor and cause of pneumonia³⁷⁻
136 ⁴⁰. Most industries are located in the southern parts of Bhutan where air pollution is expected
137 to be higher as compared to other districts. This was reflected by these sub-districts having
138 higher SMR for pneumonia. Additionally, traditional methods of cooking in rural Bhutan using
139 fire wood could also contribute to respiratory illness such as pneumonia^{41,42}.

140 The incidence of Pneumonia tends to be higher during the rainy season⁴³⁻⁴⁵. Rainfall may
141 trigger socio-ecological behavioural changes such as increased contact between people and the
142 distribution of pathogens. Further, heavy rainfall during the monsoon is likely to pollute
143 drinking water, particularly the surface water from streams, which is the main drinking water
144 source for rural populations⁴⁶. Unsafe drinking water and sanitation are important drivers of
145 pneumonia⁴⁷. Relative humidity was associated with a decrease in pneumonia incidence in this
146 study which is in concordance with other studies^{35,48}. Higher relative humidity decreases the
147 survival of lipid-enveloped viruses such as influenza A, influenza b and Respiratory Syncytial
148 Virus^{49,50}.

149 There are a number of limitations that need to be considered when interpreting the results of
150 this study. First, the study used routine case reports to measure incidence of pneumonia. Known
151 issues exist surrounding completeness and representativeness of such data. Secondly, the causal
152 organisms of pneumonia were not available and the association could be different based on the
153 organisms. Thirdly, there was no reconciliation to accommodate different levels of aggregation
154 of the climate variables (district) and the disease data (sub-district), and the climate conditions
155 were assumed to be homogeneous within a district. Lastly, unaccounted risk modifiers were
156 not included in the modelling due to a lack of available data. These important unmeasured
157 factors, such as immunization coverage, air pollution level, living standards and socio-
158 economic status, crowding, smoking, access to safe drinking water and latrine usage might
159 have resulted in confounding, which was not able to be quantified^{39,51,52}.

160 Despite these limitations, the strengths of this study are the capacity to implement the spatial
161 analysis at a relatively fine resolution, being the sub-district level, and over a long time series
162 (108 months). Traditionally, spatial patterns of infectious disease risk have been displayed at
163 larger geographical units, such as a district, province, national, regional, and global
164 scales^{46,53,54}. Such low resolution can mask localized disease patterns due to averaging⁵⁵.

165 **Conclusion**

166 Pneumonia is an important childhood disease and the introduction of pneumococcal conjugate
167 vaccines to reduce the burden of this disease is timely. Pneumonia was highly seasonal and
168 spatially heterogeneous across sub-districts. Seasonality can be explained by climatic factors
169 including temperature, rainfall and relative humidity. The spatial and temporal variability of
170 pneumonia should inform in better targeting of its prevention and control in the country through
171 rational decision making and proper resources allocation.

172 **Materials and methods**

173 **Study area**

174 Bhutan located in the Eastern Himalayas, borders China in the north and India in the east, south
175 and west. The country is divided administratively into 20 districts and 205 sub-districts, with a
176 total projected population of 741,672 in 2019 ⁵⁶. Around 62.2% (452,178) of the population
177 live in rural areas and practice subsistence farming. The altitude ranges from 75m above sea
178 level in the south to more than 7000m in the Himalayas (**Fig. 1**).

179 **Study design and data source**

180 This is a retrospective study using secondary data on pneumonia from January 2010 to
181 December 2018, stratified by sex and age (> 5 years and ≥ 5 years) at the sub-district level. The
182 data were obtained from the National Acute Respiratory Infections surveillance system, hosted
183 by the Bhutan Health Information and Management Systems (HIMS) under the Bhutan
184 Ministry of Health. These data contain all pneumonia cases treated by health centres including
185 hospitals and primary health care facilities and reported to the HIMS every month. Pneumonia
186 is defined as “ a patient with history of cough or reported breathing difficulty, and increased
187 respiratory rate (RR) or chest indrawing (RR ≥ 50 breaths per minute in children aged two
188 months or more and less than 12 months or RR ≥ 40 breaths per minute in children aged 12

189 months or more and less than 60 months⁵⁷. Daily climatic variables (rainfall, relative
190 humidity, minimum and maximum temperature) were obtained from the National Centre for
191 Hydrology and Meteorology under the Ministry of Economic Affairs of Bhutan. Monthly
192 average climatic variables were calculated for this study. Population estimates used in the study
193 were obtained from the National Statistical Bureau, Bhutan⁵⁸. Administrative boundary maps
194 were downloaded from the DIVA-GIS website⁵⁹.

195 **Crude standardized morbidity ratios**

196 An initial descriptive analysis of pneumonia incidence across the country was conducted.
197 Crude SMR for each sub-district were calculated using the following formula:

$$198 \quad Y_i = \frac{O_i}{E_i}$$

199 Where Y is the overall SMR in sub-district i , O is the total number of observed pneumonia
200 cases over the entire study period in the sub-district and E is the expected number of pneumonia
201 cases in the sub-district across the study period. The expected number was calculated by
202 multiplying the national incidence by the average population for each sub-district over the
203 study period.

204 **Exploration of seasonal patterns and inter-annual patterns**

205 The time series of pneumonia incidence was decomposed using STL weighted regression to
206 show: the seasonal pattern, inter-annual patterns and the residual variability. The STL model
207 was structured as follows:

$$208 \quad Y_t = S_t + T_t + R_t$$

209 where Y_t represents numbers of local pneumonia cases with logarithmic transformation, S_t is
210 the additive seasonal component, T_t is the trend, and R_t is the “remainder component”; t is time
211 in months^{60,61}.

212 **Spatio-temporal model**

213 A Bayesian statistical framework was deployed for spatial analysis. It provides a convenient
214 framework for the simultaneous inclusion of covariates and spatial autocorrelation in a single
215 model, while providing robust evaluation of and expression of uncertainty. The posterior
216 distributions can be used to quantify uncertainties in parameters of interest (e.g., covariate
217 effects and spatial patterns of disease risk)⁶².

218 Initially, a preliminary bivariate Poisson regression of pneumonia cases was undertaken to
219 select the covariates. The covariates with a p -value of <0.05 and the lowest Akaike's
220 information criterion (AIC) were selected. The co-linearity of the selected climatic and
221 environmental variables was tested using variance inflation factors (VIF). In the final model,
222 rainfall, maximum temperature and relative humidity were included.

223 Of the 88,560 observations stratified by sub-districts, <5 and ≥ 5 years and sex over 108 months,
224 there were 55,975 (63.2%) zero counts of pneumonia. Therefore, Zero-inflated Poisson (ZIP)
225 regression was constructed in a Bayesian framework. The first model (Model I), assumed that
226 spatial autocorrelation was not present in the relative risk of pneumonia. This model was
227 developed with selected climatic factors (rainfall, maximum temperature and relative
228 humidity), age (<5 and ≥ 5 years) and gender as explanatory variables, and an unstructured
229 random effect for sub-districts; the second model (Model II) contained a spatially structured
230 random effect in addition to the covariates; and the final model (Model III), a convolution
231 model, contained all of the components of the preceding two models. The best model with the
232 lowest DIC was selected as the final explanatory model.

233 Model III assumed that the observed counts of pneumonia, Y , for i^{th} sub-district ($i=1..205$) in
 234 the j^{th} month (January 2010-December 2018) followed a Poisson distribution with mean (μ_{ij}),
 235 that is,

$$236 \quad P(Y_{ij} = y_{ij}) = \begin{cases} \omega + 1 (1 - \omega)e^{-\mu}, & y_{ij} = 0 \\ (1 - \omega)e^{-\mu} \mu_{ij}^{y_{ij}} / y_{ij}!, & y_{ij} > 0; \end{cases}$$

$$237 \quad Y_{ij} \sim \text{Poisson}(\mu_{ij})$$

$$238 \quad \log(\mu_{ij}) = \log(E_{ij}) + \theta_{ij}$$

$$239 \quad \theta_{ij} = \alpha + \beta_1 \times \text{Age} + \beta_2 \times \text{Sex} + \beta_3 \times \text{trend}_j + \beta_4 \times \text{Rainfall}_{ij} + \beta_5 \times \text{Humidity}_{ij} +$$

$$240 \quad \beta_6 \times \text{Tempmax}_{ij} + u_i + s_i + w_i$$

241 where expected number of cases in sub-district i , month j (acting as an offset to control for
 242 population size) was represented by E_{ij} and θ_{ij} is the mean log relative risk (RR). The intercept
 243 (α), and coefficients for age (≥ 5 as reference), sex (male as reference), monthly trend, rainfall,
 244 relative humidity and maximum temperature are $\beta_1, \beta_2, \beta_3, \beta_4, \beta_5$ and β_6 . The spatially
 245 unstructured and structured random effects are represented as u_i and s_i , respectively, with u_i
 246 excluded from Model II and s_i excluded from Model I. Spatiotemporal random effect with a
 247 mean of zero and variance of σ_w^2 was denoted by w_i as in other studies^{63,64}.

248 A conditional autoregressive (CAR) prior structure was used to model the spatially structured
 249 random effect. Spatial relationships between the sub-districts were based on a ‘queen’
 250 contiguity matrix. A weight of 1 was assigned to sub-districts sharing a border and 0 otherwise.
 251 A flat prior distribution was specified for the intercept, whereas a non-informative normal prior
 252 distribution was used for the coefficients. The priors for the precision of unstructured and
 253 spatially structured random effects were specified using non-informative gamma distributions
 254 with shape and scale parameters equal to 0.01.

255 The model was run for an initial 10,000 iterations, which were then discarded. Subsequently,
256 visual inspection of posterior density and history plots were used to note convergence at
257 intervals of 20,000 iterations. Convergence occurred at approximately 100,000 iterations for
258 all models. Following convergence, posterior distributions from model parameters were stored
259 for inference. Markov Chain Monte Carlo simulation was used to estimate model parameters
260 ⁶⁵. Summaries of parameters were calculated, including posterior mean and 95% credible CrI.
261 In all analyses, an α -level of 0.05 was adopted to indicate statistical significance (as indicated
262 by 95% CrI for relative risks (RR) that excluded 1).

263 Seasonality decomposition was carried out using the R statistical package, release 3.3.1. The
264 ZIP regression model was constructed using WinBUGS software, version 1.4.3 (MRC
265 Biostatistics Unit 2008)⁶⁶. ArcMap 10.5 software (ESRI, Redlands, CA) was used to generate
266 maps of the posterior means of the unstructured and structured random effects and the
267 spatiotemporal random effects.

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446 **Authors contribution**

447 KW and KP were involved in the conception of the study. KW and TT undertook the analysis.
448 KP and CT obtained the data. KW and KP drafted the manuscript. ACAC, PG and DJG

449 critically reviewed and edited the manuscript. All authors read and approved the final
450 manuscript.

451 **Competing interest**

452 Authors declare there is no competing interest.

453 **Ethical approval and patient confidentiality**

454 Administrative approval to use these datasets was provided by the Ministry of Health, Bhutan.

455 This study was a low-risk study since the surveillance data did not contain identifying
456 information on individual participants.

457 **Data availability**

458 The datasets of this current study will be made available from the corresponding author on
459 reasonable request.

460

461 **Figures**

462 **Figure 1 Map of Bhutan with districts and sub-districts with altitude.**

463 **Figure 2 Decomposed monthly cases of pneumonia: (a) under 5 years and (b) 5 years**
464 **and older during the study period, 2010-2018.**

465 **Figure 3 Crude standardized morbidity ratios (SMR) of pneumonia by sub-districts the**
466 **study period, 2010-2018.**

467 **Figure 4 (a) Spatial distribution (b) significance map of the posterior means of**
468 **unstructured random effects of pneumonia in Bhutan, 2010-2018.**

469 **Figure 5 Trend of pneumonia by sub-districts of Bhutan during the study period, 2010-**
470 **2018.**

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474 **Table 1** Yearly incidence of pneumonia stratified by age.
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Year	Under 5 years			5 years and older		
	Cases	Population	Incidence*	Cases	Population	Incidence*
2010	9,204	77,161	119.28	3,369	70,582	47.73
2011	7,975	78,618	101.44	3,210	718,702	4.47
2012	9,939	80,985	122.73	3,683	741,572	4.97
2013	8,956	81,899	109.35	3,064	749,937	4.09
2014	9,434	82,947	113.74	3,669	759,536	4.83
2015	7,489	84,009	89.15	3,037	769,258	3.95
2016	8,150	85,084	95.79	3,023	779,105	3.88
2017	5,883	86,173	68.27	2,606	789,077	3.30
2018	4,777	87,276	54.73	2,547	799,177	3.19

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477 *incidence per 1,000 population

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479 **Table 2** Regression coefficients, relative risk and 95% credible interval from Bayesian
480 spatial and non-spatial models of pneumonia cases in Bhutan, January 2010-December
481 2018.

Model/Variable	Coeff, posterior mean (95% CrI)	RR, posterior mean (95% CrI)
Model I (Unstructured)		
α (Intercept) [†]	-4.18 (-4.32, -4.13)	
Age (base over 5 years)	3.06 (3.04, 3.07)	21.26 (20.95, 21.59)
Sex (base male)	0.08 (0.06, 0.09)	1.08 (1.066, 1.094)
Mean monthly trend	-0.12 (-0.14, -0.10)	0.886 (0.870, 0.902)
Rainfall (10mm)	0.07 (0.01, 0.13)	1.072 (1.007, 1.140)
Relative humidity**	-0.09 (-0.10, -0.08)	0.914 (0.903, 0.925)
Maximum temperature (°C)	0.25 (0.24, 0.26)	1.286 (1.272, 1.301)
Probability of extra zero	0.26 (0.21, 0.30)	
Heterogeneity		
Unstructured	0.43 (0.35, 0.53)	
Structured (trend)	1.82 (1.43, 2.29)	
DIC*	206,040	
Model II (Structured)		
α (Intercept) [†]	-4.18 (-4.32, -4.13)	
Age (base over 5 years)	3.06 (3.02, 3.09)	21.26 (20.55, 22.02)
Sex (base male)	0.08 (0.05, 0.10)	1.08 (1.052, 1.108)
Mean monthly trend	-0.12 (-0.14, -0.10)	0.886 (0.869, 0.903)
Rainfall (10mm)	0.07 (-0.01, 0.14)	1.070 (0.999, 1.014)
Relative humidity**	-0.09 (-0.10, -0.08)	0.914 (0.902, 0.927)
Maximum temperature (°C)	0.25 (0.24, 0.27)	1.287 (1.270, 1.304)
Probability of extra zero	0.18 (0.17, 0.19)	
Heterogeneity		
Structured (spatial)	0.09 (0.07, 0.11)	

Structured (trend)	1.82 (1.42, 2.28)	
DIC	206,093	
Model III (Mixed)		
α (Intercept) [†]	-4.15 (-4.36 -3.91)	
Age (base over 5 years)	3.06 (3.04, 3.07)	21.26 (20.72, 21.82)
Sex (base male)	0.08 (0.06, 0.09)	1.080 (1.059, 1.101)
Mean monthly trend	-0.12 (-0.14, -0.10)	0.886 (0.870, 0.902)
Rainfall (10mm)	0.07 (0.01, 0.13)	1.071 (1.000, 1.014)
Relative humidity**	-0.09 (-0.10, -0.08)	0.914 (0.903, 0.926)
Maximum temperature (°C)	0.25 (0.24, 0.26)	1.287 (1.271, 1.303)
Probability of extra zero	1.201 (1.191, 1.211)	
Heterogeneity	0.60 (0.42, 1.02)	
Unstructured	1.68 (0.13, 8.04)	
Structured (spatial)	1.68 (0.13, 8.04)	
Structured (trend)	1.82 (1.42, 2.28)	
DIC	206,058	

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* best-fit model; ** Lagged three months, [†]coefficient

Abbreviations: coeff-coefficients; CrI- credible interval; RR-relative risk; DIC- deviation information criterion