

Joint Modeling of Hypertension Measurements and Time-to-Onset of Preeclampsia Among Pregnant Women Attending Antenatal Care Service at Arerti Primary Hospital, North Shoa, Ethiopia

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ABSTRACT

Background: Preeclampsia is a hypertensive disorder of pregnancy that affects 2-8% of pregnant women. It is the major cause of maternal and perinatal morbidity and mortality worldwide. The purpose of this study was to identify factors associated with hypertension measurements and time-to-onset of preeclampsia among pregnant women attending antenatal care service at Arerti Primary Hospital.

Methodology: A retrospective longitudinal study design was employed on a total of 201 pregnant women attending the antenatal clinic of Arerti Primary Hospital between September 2018 and June 2019. A closed-form sample size formula for estimating the effect of the longitudinal data on time-to-event was used. To analyze our data we employed descriptive method, linear mixed effect model, Cox-PH model and joint models for longitudinal and survival outcomes. Relevant demographic and clinical covariates were included in sub models.

Results: This study revealed that baseline age, visiting times, weight, diabetes, history of PE and parity had significantly associated with mean change in the BP measurements. From the Cox model result, age, weight, history of PE and marital status were associated with a significant hazard of developing preeclampsia. The univariate joint models reveal that the each longitudinal

BP measurements are significantly associated with hazard of developing preeclampsia. From the bi-variate joint model; only DBP is significantly associated with risk of developing PE.

Conclusion: As the result obtained in this study, we summarized that, age, weight, history of PE and marital status had a significant effect on time to developing preeclampsia. Furthermore, due to significance of association between the longitudinal BP measurements and time to onset of preeclampsia, joint model analysis was suggested as it incorporates all information simultaneously and provides valid and efficient inferences over separate models analysis.

Key words: Bi-variate joint model, Diastolic Blood Pressure, Longitudinal sub-model, Systolic Blood Pressure, Survival sub-model, Time to onset of preeclampsia, univariate joint model

Plain English Summary

Preeclampsia is a leading cause of maternal mortality and morbidity worldwide. According to Ethiopian Demographic Health Survey, (2016), the estimate of the maternal mortality ratio is 412 deaths per 100,000 live births; that is, for every 1000 births, there are about 4 maternal deaths. Three outcome variables were considered for this study; the longitudinal Systolic Blood Pressure (SBP), Diastolic Blood Pressure (DBP) in units of millimeters of mercury (mmHg) and survival outcome, time to the development of preeclampsia or delivery from antenatal care service follow up of pregnant women. All Pregnant women on antenatal care service at the booking visit who have two and more than two visits were extracted and followed up until delivery or development of the outcome. The longitudinal SBP and DBP measurements were measured at the following 4 time points over a 10-month time period; first visit (4th- 13th weeks gestation), second visit (14th - 27th weeks gestation), third visit (28th-35th weeks gestation) and last visit (36th-40th weeks gestation). For this study we draw a sample of 244 pregnant women following antenatal care service from a population of 1220 pregnant women following antenatal care service by applying systematic random sampling. In conclusion the finding from the analysis of longitudinal responses (SBP and DBP) and survival time to onset of preeclampsia of pregnant women could be important for pregnant women, health programmers, antenatal care service providers and future researchers in prevention and management of preeclampsia.

Background

Preeclampsia (PE) is a hypertensive disorder of pregnancy that affects 2-8% of pregnant women worldwide and characterized by elevated blood pressure and protein in the urine during second or third trimester of gestation (Ghulmiyyah and Sibai, 2012). It most commonly develops during the last trimester, but it can happen at any time in the second half of pregnancy, during labor, or even up to six weeks after delivery. It is a serious and poorly understood complication of pregnancy, which can progress to eclampsia and health problems among mothers and their babies (Choudhury and Friedman, 2012).

The world health organization, (2011) and American College of Obstetrics and Gynecology, (2019) develop guidelines for the prevention, early diagnosis, and management of PE and other Hypertensive Disorders of pregnancy. Low dose aspirin and calcium are recommended for prevention (Bodnar, *et al.*, 2007 ; Rumbold, *et al.*, 2008). Despite such recommendations, delay in identification and diagnosis of PE, the prevalence of preeclampsia occurs in 10% of pregnancies in African women, which is significantly higher than the global average of approximately 2% (Nakimuli, *et al.*, 2014).

According to Ethiopian Demographic Health Survey, (2016), the estimate of the maternal mortality ratio is 412 deaths per 100,000 live births; that is, for every 1000 births, there are about 4 maternal deaths. The Ethiopian National Emergency Obstetric and Newborn Care (EMONC) showed that preeclampsia contributed for the complication of approximately 1% of all deliveries and 5% of all pregnancies. Moreover, 16% of direct maternal mortality was due to preeclampsia/eclampsia (Gaym, *et al.*, 2011). Ethiopia is one of the five countries that together account for 50 percent of the world's maternal deaths and rural areas mostly contribute to maternal deaths (Gaym, *et al.*, 2011). A maternal mortality trend analysis showed an increasing trend of preeclampsia in Ethiopia.

Several studies have been conducted related to pregnant women to determine risk factors associated with survival time and longitudinal outcomes separately and epidemiological studies for example (Mohammednur and Abdo, 2019; Mo *et al.*, 2019 ; Musa *et al.*, 2018 ; Tessema *et al.*, 2015; Vata *et al.*, 2015; Gaym *et al.*, 2011; Garomssa and Dwivedi, 2008; Jacobs *et al.*, 2003). However, the linear mixed model for longitudinal data and the Cox proportional hazards

model for time-to-event data do not consider dependencies or interrelationships between these two different data types (longitudinal and time-to-event data) separately (Yanga *et al.*, 2016). Hence, appropriate statistical method (joint model) was proposed by several researchers in order to identify potential relationships between the longitudinal measures and disease risk (Rizopoulos, 2012; Wulfsohn and Tsiatis, 1997; Faucett and Thomas, 1996). Joint modeling of longitudinal and survival data is preferable than separate analysis as it incorporates all information simultaneously and provides valid and efficient inferences for the given data (Yanga *et al.*, 2016). Hence, the aim of this study focused on jointly modeling of hypertension measurements and time-to-onset of preeclampsia among pregnant women attending antenatal care service at Arerti Primary Hospital.

Methods

Study Area and Population of Interest

The study was conducted in Arerti Primary Hospital located in Arerti town in Minjar Shenkora Woreda in the North Shoa Zone at Amhara regional state of Ethiopia. The study population covered of all pregnant women attending antenatal care service at Arerti Primary Hospital between September 2018 and June 2019.

Study Design

A retrospective longitudinal study design was conducted at the antenatal clinic of Arerti Primary Hospital from September 2018 and June 2019. Both the longitudinal and survival data was extracted from the pregnant woman's follow up cards which contains socio-demographic characteristics, baseline clinical, laboratory measurement and other information of all pregnant women attending antenatal care service.

Study Population

A total of 201 of pregnant women were included from women's pregnant at ANC follow up at Arerti primary hospital. The Sample size (number of events) determination formula for studying the relationship between event time and the longitudinal process was calculated using the formula:

$$d = \frac{2(Z_{\alpha} + Z_{\beta})^2}{\sigma_s^2 \alpha^2}$$

And the required total number of subjects (pregnant women) can be computed as:

$$n = d / P(\text{event})$$

(Chen *et al.*, 2011). The parameters were estimated from pilot survey because no previous joint analyses are available to determine appropriate hypothesized values of association parameters. The assumptions include 95% Confidence level and 80% power. Thus the final sample size was 201. A systematic random sampling method was adopted for selecting a representative sample from the list of the medical charts that contain the list of pregnant women name and identification number. A total of 1220 pregnant women were at ANC follow up during the study period. K was calculated by dividing the total number of pregnant women at ANC followup in the hospital in the study period by the sample size ($K = 1220/244 = 5$). Every 5th pregnant women at ANC followup in the hospital during the study period was selected for assessment.

Inclusion and Exclusion Criteria

All Pregnant women on ANC followup service at the booking visit who have two and more than two visits were extracted and followed up until delivery or development of the outcome included to the study. Those pregnant women with significant proteinuria at booking and blood pressure (BP) of $\geq 140/90$ millimeter mercury (mmHg) readings at ≤ 20 weeks gestation were excluded from the study.

Response Variable of the Study

Three outcome variables were considered for this study; the longitudinal Systolic Blood Pressure (SBP) , Diastolic Blood Pressure (DBP) in units of millimeters of mercury (mmHg) and survival outcome; time to the development of preeclampsia or delivery from ANC follow up of pregnant women

Data Analysis Methods

separate and joint methods of longitudinal data analysis and survival data analysis were fitted to answer the research objectives. The analysis consists; exploratory data analysis, linear mixed

effects model for the longitudinal data, Cox proportional hazards model for the time to event data and joint model for both longitudinal and time to event data.

Results

Descriptive Statistics

During the 10 months study period (September 2018 to June, 2019), a total of 201 of pregnant women who attended antenatal care (ANC) are obtained in Arerti Primary Hospital. In the supporting file (Table 4.1) descriptive statistic of baseline covariate of those women was illustrated.

We observed 39 out of the 201 normotensive women included in the study developed PE and 162 of them were censored observations during the study period. The mean age and weight of the study pregnant women were 26.0 and 58.2, respectively. Out of the total sample, 23 (11.4%) had diabetes mellitus. Regarding the previous history of preeclampsia, 172 (85.6%) had no previous history of Preeclampsia. Out of the total pregnant women who examined, 177(88.1%) of the women had no consecutive previous abortion.

Out of the total number of the current pregnancy, 172(85.6%) were singleton pregnancy and 29 (14.4%) of them were twin; and out of singleton pregnant women 154 (95.1%) were censored, the remaining 18 (46.2%) were event occurred; and out of twin pregnant women 8 (4.9%) were censored, the remaining 21 (53.8%) were event occurred. Majority of the pregnant women, 175(87.1%) of them were married and 26 (12.9%) were unmarried; and out of married women, 159 (98.1%) were censored, the remaining 16 (41.0%) women were event occurred; and out of unmarried women 3 (1.9%) were censored, the remaining 23 (59.0%) were event occurred.

Before fitting the univariate and bi-variate linear mixed-effects models, first we should checked the assumption of normality of the data. To check normality assumption of linear mixed effects model, normal Q-Q plots of the BP measurements with corresponding Shapiro-Wilk tests of normality were checked. In the supporting file Figure 4.1a and 4.1b shows the normal Q-Q plots of the SBP and DBP normality of the distributions of BP measurements. Similarly result from Shapiro-Wilk test are not significant for SBP and DBP implying that the BP data set appear to follow a normal distribution(see supporting information file Table 4.2).

Summaries of Longitudinal blood pressures (BP)

The bi-variate continuous longitudinal SBP and DBP measurements were measured at the 4 time points over a 10-month time period; first visit (4th- 13th weeks gestation), second visit (14th - 27th weeks gestation), third visit (28th-35th weeks gestation) and last visit (36th-40th weeks gestation).

A total of 201 pregnant women were observed at first visit and second visit. Then, 165(82.1%) pregnant women were observed at the third visit and 101(50.2%) pregnant women stayed at the last visit during follow-up periods. Mean with the corresponding standard deviation at each time points with respective sample sizes for both outcome variables were shown in Table 4.3. The numbers of participants are decreasing in successive time points. The average SBP and DBP were 116.84 and 66.99 at first visit respectively. There was a general increment in the mean value up to the last time point. When we look at the standard deviations, smaller variations were observed at first visit and higher variations were seen at Second visit.

Table 4.3: Summary Measures of the Two Longitudinal Responses at Each Time Points

	Time	First visit	Second visit	Third visit	Fourth visit
SBP	Sample points (%)	201(100%)	201(100%)	165(82.1%)	101(50.2%)
	Mean	116.84	122.88	126.25	128.89
	SD	10.63	11.71	10.27	9.23
	Time	Frist visit	Second visit	Third visit	Fourth visit
DBP	Sample points (%)	201(00%)	201(100%)	165(82.1%)	101(50.2%)
	Mean	66.99	74.81	79.81	81.45
	SD	8.49	10.46	9.51	8.73

Univariate Linear Mixed Effect Model Result

The separate longitudinal analysis was started from the fixed effect modeling to select appropriate covariates that predicts SBP and DBP. To select the fixed effect components of the response variable SBP and DBP, including all covariates without considering the corresponding different random effects using a backward variable selection method. The predictor variables visit, age, weight, multiplicity, PE, gravidity, parity and diabetes were considered for linear mixed

model for longitudinal BP. To identify the appropriate co-variance structure for the two longitudinal response variables; the three commonly used co-variance structures were considered. In the supporting file Table 4.4 shows that the First Order Auto-regressive (AR (1)) co-variance structure best fits for our data compared to the other co-variance structures.

In order to retain the random effects from the model, it is better to fit the linear mixed effects model with different random effects. In the supporting file (Table 4.5) shows that random intercept and random slope model were used in linear mixed effect model that appropriately predicts the mean change of SBP and DBP over time.

Table 4.6 shows the result of linear mixed effect model of SBP and DBP which found the covariates visit, age, weight, PE, multiplicity, parity and visit, age, weight, PE, multiplicity, parity, diabetes were significantly associated with mean change in the SBP and DBP at 5% level of significance respectability.

Table 4.6: Parameter Estimate of the Univariate Linear Mixed Effect Models of SBP and DBP

Variables	SBP				DBP			
	Estimates	Std.Error	T	P	Estimates	Std.Error	T	P
Intercept	118.350	5.77	20.47	<.0001	65.790	4.73	13.86	<.0001
Visit	0.50	0.081	12.71	<.0001	0.625	0.028	21.21	<.0001
Age	0.538	0.1415	3.81	0.0002	0.534	0.115	4.63	<.0001
Weight	0.1404	0.053	2.57	0.0101	0.125	0.043	2.84	0.0045
PE (No)	Ref.							
Yes	5.010	2.356	2.14	0.0324	6.751	1.948	3.40	0.0007
Multiplicity (Singleton)	Ref.							
Twin	6.012	2.439	2.46	0.0140	5.746	2.003	2.81	0.0051
Gravidity(≥2)	Ref.							
1	2.166	2.101	1.03	0.3031	0.692	1.783	0.40	0.689
Parity(≥2 deliveries)	Ref.							
1 delivery	3.406	1.222	2.79	0.0055	3.049	1.000	3.04	0.0024

Nulliparous	-5.185	3.162	-1.64	0.1017	-5.840	2.632	-1.64	0.1017
Diabetes(No)	Ref.							
Yes	3.494	2.484	1.41	0.1602	5.459	2.184	2.51	0.0121

Bi-variate Linear Mixed-Effects Analysis of SBP and DBP

The parameter estimates of the final fitted bi-variate linear mixed model are presented in Table 4.7. The variables visit, age, weight, multiplicity and parity were found significant at 5 percent level of significance for the change in SBP and DBP. However, PE and diabetes were significantly associated with mean change in the DBP at 5% level of significance. Gravidity was the only variable did not significant in both responses. The intercept 65.7 and 117.79 with standard error of 4.81 and 5.87 represent estimates of the average level of DBP and SBP during the first follow up time respectively.

A parameter estimate for weight indicates a one-kg increase in weight was associated with a normal increase of 0.129 mmHg (Std error =0.044) in DBP and 0.144 mmHg (Std error = 0.054) in SBP respectively. A parameter estimate of age indicates a one year increase in age was associated with a normal increase of 0.46 mmHg (Std error = 0.114) in DBP and 0.57 mmHg (Std error = 0.14) in SBP respectively. The average intercepts for multiplicity (twin) are 6.311 and 6.24, which indicates that on average pregnant women with twin pregnancy started with the higher DBP and SBP measure than singleton pregnancy at baseline.

Table 4.7: Bi-variate Mixed Model fixed Parameter Estimates, Standard Error and P -Values for the final Model

Variables	DBP				SBP			
	Estimate	Std.Error	Z	P	Estimate	Std.Error	Z	P
intercept	65.709	4.81	13.64	<.0001	117.790	5.87	20.04	<.0001
Visit	0.640	0.029	21.79	<.0001	0.530	0.030	17.44	<.0001
Age	0.460	0.114	4.05	<.0001	0.570	0.148	4.07	0.0001
Weight	0.129	0.044	2.92	0.0037	0.144	0.054	2.67	0.0079
PE(No)	Ref.				Ref.			
Yes	6.487	2.005	3.24	0.0013	4.573	2.3971	1.91	0.057

Multiplicity (Singleton)	Ref.				Ref.			
Twin	6.241	2.042	3.06	0.0024	6.311	2.4671	2.56	0.0108
Diabetes(No)	Ref.				Ref.			
Yes	5.606	2.2434	2.5	0.0128	4.121	2.536	1.62	0.1048
Gravidity(≥ 2)	Ref.				Ref.			
1	0.652	1.8207	0.36	0.720	1.908	2.141	0.89	0.373
Parity(≥ 2 deliveries)	Ref.				Ref.			
Nulliparous	-6.417	2.687	-2.39	0.0173	-5.709	3.210	-1.78	0.0759
1 delivery	3.002	1.0182	2.95	0.0033	3.343	1.235	2.71	0.007
Residual	36.911	3.28	11.24	<0.0001	34.77	2.83	12.25	<0.0001

Survival Data Analysis Result

The pregnant women were followed up for the total of 40 weeks. The median survival time was found to be 31 weeks. At the end of the follow up, 39 (19.4%) women developed preeclampsia, the remaining 162 (80.6%) didn't developed preeclampsia and loss to follow-up.

Separate graphs of the estimates of the Kaplan-Meier survivor functions for the categorical covariates are presented in order to see the patterns of survival experience those different covariates. The survivor ship pattern of one is lying above another means the group defined by the upper curve has a better survival than the group defined by the lower curve(see the supporting file Figure 4.2).

The Log-rank test was performed to investigate the significance of the observed difference in the Kaplan-Meier estimates of the survivor functions among different categories of the factors. The result indicates that there were significant differences in survival probability of pregnant women in covariates of marital status, abortion, diabetes, PE, substance, multiplicity, gravidity and parity at 5% level of significance (see the supporting file Table 4.8).

Cox Proportional Hazard Model

In Cox proportional regression model (see supporting file Table 4.9), the factors associated with a significant hazard of developing preeclampsia during the gestational follow up time were age, weight, PE, multiplicity and marital status. Looking for the age of pregnant women, the risk of

developing preeclampsia increased by 43.3% (HR = 1.43, Std. Error = 0.074, P – value < 0.0001). The hazard of developing preeclampsia for those older age of pregnant women were 43.3 % higher risked than those are younger women. Controlling for other risk factors, the risk of developing preeclampsia, as estimated from a Cox model, was HR = 1.188 times higher per unit difference in weight. That is, holding the other covariates constant, a higher value of weight is associated with a poor survival. Regarding multiple pregnancy of women, twin pregnant had higher risk of developing preeclampsia compared to singleton pregnant (HR = 2.633, P – value = 0.0307). Women with PE had 1.07 times greater risk of developing PE compared to those with no prior history (HR = 1.074, Std. Error = 0.76570, P – value = 0.0004). The estimated risks of preeclampsia for unmarried pregnant women compared to those married women were 2.66 (HR = 2.66, Std. Error = 0.54593, P – value = 0.0003). This means the hazard rate of preeclampsia for unmarried pregnant women is 2.66 times more likely to develop preeclampsia than married pregnant women.

Results for Univariate Joint Analysis

The joint model of the BP measurements and time to preeclampsia was analyzed by using R software, JM package. The two separate joint models showed that both SBP and DBP were significantly associated with preeclampsia risk. Table 4.10, revealed that most covariates found to be statistically significant on joint analysis and a strong association between the longitudinal bio-markers (BP). **For longitudinal sub models**, the covariates age, weight, multiplicity, parity and PE are found to be significantly associated the longitudinal change of BP measurements. A one unit increase in age was associated with 0.53 and 0.54 increased in the mean DBP and SBP respectively. The mean changes in DBP and in SBP were 5.87 and 2.439 times higher for twin pregnant when compared to singleton. Comparing mean change in by PE, the mean change in DBP and in SBP were 6.751 and 5.01, respectively times higher for women with PE compared to without PE.

The survival sub models revealed that baseline maternal age, weight, PE and marital status were significant in both sub-models. But multiplicity was significant in survival sub models of SBP only. The association parameters (α_k) were significantly different from zero (p – value < 0.0001) indicating that there is strong association between both DBP and SBP and the risk for

preeclampsia. The positive value of the association parameters ($\alpha_1=0.046$ and $\alpha_2=0.085$) indicated that true value of DBP and SBP were positively associated with the risk of preeclampsia, and with unit increase in the longitudinal DBP and SBP the risk of preeclampsia were increased by $\exp(0.046) = 1.05$ and $\exp(0.085) = 1.09$ respectively.

Table 4.10: Results of Univariate Joint Modeling of Time to Preeclampsia Event and Longitudinal BP.

Longitudinal DBP Sub Model					Longitudinal SBP Sub Model			
Variables	Est.	Std.Error	T	P	Est.	Std.Error	T	P
Intercept	65.80	4.74	13.87	<.0001	118.36	5.78	20.48	<.0001
Visit	0.626	0.029	21.22	<.0001	0.51	0.082	12.72	<.0001
Age	0.534	0.115	4.63	<.0001	0.538	0.1415	3.81	0.0002
Weight	0.126	0.044	2.85	0.0046	0.1405	0.054	2.58	0.0101
PE (No)	Ref.				Ref.			
Yes	6.751	1.949	3.46	0.0006	5.009	2.356	2.13	0.0340
Multiplicity (Singleton)	Ref.				Ref.			
Twin	5.847	2.004	2.92	0.0037	6.012	2.439	2.46	0.0140
Gravidity(≥ 2)	Ref.				Ref.			
1	0.692	1.783	0.39	0.6978	2.166	2.101	1.03	0.3031
Parity(≥ 2)	Ref.				Ref.			
1delivery	3.049	1.000	3.05	0.0024	3.406	1.222	2.79	0.0055
Nulliparous	-5.84	2.632	-1.64	0.1017	-5.185	3.162	-1.64	0.1017

Diabetes(No)	Ref.				Ref.			
Yes	5.459	2.184	2.50	0.0128	3.494	2.484	1.41	0.1602
Survival Sub-Model					Survival Sub-Model			
Age	0.309	0.069	4.436	<0.001	0.380	0.061	6.180	<0.0001
Weight	0.158	0.032	5.97	0.001	0.138	0.022	6.066	<0.0001
PE(No)	Ref.				Ref.			
Yes	0.997	0.40	2.45	0.0142	0.697	0.423	1.646	0.0496
Multiplicity (Singleton)	Ref.				Ref.			
Twin	0.201	0.638	0.314	0.07	1.306	0.671	1.944	0.048
Martial(Mar ried)	Ref.				Ref.			
Unmarried	0.901	0.362	2.502	0.0001	0.86	0.392	2.19	0.0009
Parity(≥ 2)	Ref.				Ref.			
Nulliparous	-0.082	0.674	-0.123	0.9024	-1.159	0.864	-1.341	0.1798
1 delivery	-0.407	0.501	0.812	0.4164	-0.109	0.572	-0.191	0.8484
Assoct(α_k)	0.046	0.002	23.49	<0.0001	0.085	0.005	14.763	<0.0001

$k = 1,2$

Results for Bi-variate Joint Analysis

In this section we fitted the bi-variate joint model of longitudinal and survival data to analyze the two longitudinal bio-markers (DBP and SBP) simultaneously through the linear mixed effect model and the cox model setup.

Table 4.11 presents the estimation results of mean intercept, mean slope and the fixed effect in the longitudinal sub-model. According to this, we have observed a significant positive trend for DBP and SBP. The covariates visiting time, age, weight, PE and parity were appeared to be associated with the change in SBP and visiting time, age, weight, PE, parity and diabetes were significant factors for the change in DBP at 5% level of significance.

Table 4.11: Parameter Estimates of Bi-variate Longitudinal Sub Model in Joint Analysis

SBP					DBP			
Variables	Estimate	Std.Error	Z	P	Estimate	Std.Error	Z	P
Intercept	117.790	5.87	20.04	<.0001	65.70	4.81	13.64	<.0001
Visit	0.530	0.030	17.44	<.0001	0.640	0.029	21.97	<.0001
Age	0.460	0.14	4.05	0.0001	-0.460	0.114	-4.05	<.0001
Weight	0.144	0.054	2.67	0.0079	0.129	0.044	2.92	0.0037
PE (No)	Ref.							
Yes	4.570	2.39	1.91	0.057	6.50	2.40	2.70	0.0013
Mul.(Singleton)	Ref.							
Twin	6.016	2.436	2.47	0.013	5.853	2.003	2.13	0.034
Gravidity(≥ 2)	Ref.							
1	2.177	2.099	1.04	0.3002	0.702	1.783	0.39	0.6941
Parity(≥ 2 deliveries)	Ref.							
1 delivery	3.409	1.228	2.78	0.0057	3.051	1.003	3.04	0.0025
Nulliparous	-5.204	3.158	-1.65	0.1	-5.856	2.631	-2.23	0.0265
Diabetes(No)	Ref.							
Yes	3.491	2.483	1.41	0.1604	5.458	2.184	2.5	0.0128

The Parameter Estimates of Survival Sub-Model is presented in Table 4.12. The association parameter γ_1 for SBP and γ_2 for DBP depicts the association between the two longitudinal responses and the event.

Table 4.12 shows observed from those association parameters; a unit increase of DBP is associated with $\exp(0.0263) = 1.026$ increases in women’s risk of developing PE assuming that the baseline covariates and SBP remain the same. But SBP was not significantly associated with hazard of PE at 5% level of significance ($\gamma_2 = 0.1458, P - \text{value} = 0.6039$) once DBP was adjusted in the model.

The fitted model also identified several other risk factors for PE, i.e. women’s age, weight; marital status and PE appeared to be risk of PE. This indicates that the hazard ratio of the baseline age of the women was $\exp(0.25) = 1.28$ means that as the age increases the hazard of PE increases.

Regarding baseline weight of women, a one unit increase in weight of a woman leads to increase the hazard of PE by $\exp(0.1694) = 1.185$. Looking at the marital status of the women, the hazard of those women who were unmarried were 47% higher than ($\exp 0.315) = 1.47, p - \text{value} = 0.0041$) those who were married.

A comparison between the separate cox model with uni-variate joint model and bi-variate joint model reveals some interesting features. Most importantly, we found that multiplicity of pregnancy is found to be significant in separate analysis of the survival outcome, but in bi-variate joint analysis and univariate joint analysis (Longitudinal DBP and time to PE) was not significant.

Table 4.12: Parameter Estimates of Survival Sub-Model in Bi-variate Joint Analysis

Variables	Estimate	Std.Error	Z	P
Age	0.250	0.065	3.85	0.0001
Weight	0.1694	0.0238	7.1221	0.0001
PE(No)	Ref.			

Yes	0.44	0.201	2.204	0.01475
Multiplicity(Singleton)	Ref.			
Twin	0.6164	1.4372	0.4289	0.6680
Parity(≥ 2 deliveries)	Ref.			
Nulliparous	0.0786	1.3225	0.0595	0.9526
1 delivery	-1.6634	1.5203	-1.0941	0.2739
Martial(Married)	Ref.			
Unmarried	0.3149	0.1547	2.093	0.0041
Assoc. (γ_1)	0.0263	0.0068	3.895	0.001
Assoc. (γ_2)	0.1458	0.2810	0.5188	0.6039

Discussion

This study attempted to assess predictors that are associated with hypertension measurements and survival time/time to preeclampsia at Arerti Primary Hospital. So as to address mentioned objectives four different analyses were explored; the bi-variate linear mixed-effects model, Cox proportional hazards model, univariate joint modeling for each longitudinal outcome and event data and bi-variate joint model of longitudinal BP and time to preeclampsia with the sets of both continuous and categorical covariates. The longitudinal process was characterized by linear mixed (Laird and Ware, 1982) sub models, and the survival process was characterized by Cox Proportional (Cox, 1972) sub model. The Maximum likelihood via EM algorithm was used to estimate unknown parameters of the joint model (Rizopoulos, 2012; Wulfsohn and Tsiatis, 1997). In the separate analysis of the longitudinal data, the assumption of normality was checked using normal Q-Q plots with corresponding Shapiro-Wilk tests. Both plots and test of the data indicate that there are no a deviation from normality assumption. This data was analyzed using different plots (exploratory data analysis) followed by model based outputs. From individual profile plots, we observed the existence of variability in BP measurements within and between individuals but,

the exploratory analysis result for the mean structure also suggested that on average, BP measurements increases in a linear pattern over time.

From the final model of LMM, we have found that visiting time, weight, parity, multiplicity of pregnancy, gestational diabetes and previous history of PE were found to be significant at 5% significance level similar to the study by (Rebelo *et al*, 2014). Our finding revealed that baseline age is an important socio-demographic predictor of BP, BP increases with increasing in baseline age; so age at baseline was a significant predictor for BP of pregnant women which is supported by study like (Mo *et al*, 2019). Results from this study demonstrate that the correlation between the two responses (DBP and SBP) estimated to be 0.96. The finding provides direct evidence that SBP and DBP measure were dropped at the beginning of pregnancy and increased throughout pregnancy and the result was supported by studies like (Mo et al, 2019).

Women with previous history of preeclampsia and baseline weight of gestation had a significantly higher risk and hazard of developing preeclampsia during the gestational follow up time (Musa *et al*.2018). Our study also supports this findings in addition to other risk factors age of pregnant women, marital status and previous history of PE of the women. In this study, the association of maternal age and risk of preeclampsia was found to be significant ($HR = 1.43, Std. Error = 0.074, P - value < 0.0001$). This is congruent with the study conducted by (Belay and Wudad, 2019; Tessema, *et al.*, 2015). This study showed marital status had statistically significant association with the development of preeclampsia in both separate and joint model. The finding was inline with the study conducted in Dessie (Tessema, *et al.*, 2015).

On the other hand, we observed the significance of the association parameters (α) between two outcomes in univariate joint modeling and the estimated association parameters are $\alpha_1=0.046$ and $\alpha_2=0.085$, $p - value < 0.0001$ indicating that the true value of DBP and SBP were positively associated with the risk of preeclampsia in a 40 weeks follow-up period and similar with the study conducted by Baschat, *et al.* (2018) who showed that average first-trimester DBP and second-trimester SBP measurements are independent risk factors for preeclampsia.

CONCLUSION

In this study separate, joint analysis of a single longitudinal outcome and bi-variate longitudinal outcomes and time to event data were conducted to establish the relationship between

longitudinal bio-marker measurements and the duration to preeclampsia. The finding of our study suggests that there are different risk factors for PE and the change of blood pressure measurements. Blood pressure measurements are influenced considerably by visit (follow-up time), age, baseline weight and multiplicity of pregnancy. The covariates age, weight, multiplicity of pregnancy, previous history of PE and marital status were found to be statistically significant effect for time to preeclampsia.

The pregnant women at ANC that their SBP and DBP measured repeatedly over time are associated with the risk for developing PE during their pregnancy, implying that women with higher baseline SBP and DBP have more likely to the event occurred; women with smaller baseline SBP and DBP have less likely to occur an event than women with higher SBP and DBP. Generally, the estimated association parameters indicate longitudinal BP and the survival time were positively associated, implies higher values of the BP associated with a worse survival. Therefore, due to significance of association between the longitudinal and survival outcomes, joint model analysis were suggested over separate models analysis as they incorporate all information simultaneously and provide valid and efficient inferences.

Declarations

Ethical and Consents

The research thesis has checked and approved by ethical clearance committee of University of Gondar, and the medical director's office of Arerti Primary Hospital granted permission to use the patients' data for this study. For the purpose of confidentiality, there were no linkages with individual patients and all data had no personal identifier and were kept confidential.

Consent for publication: Not applicable

Availability of data and materials: Yes it is applicable

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Contributions of the Authors

Dawit Baye Haile: Contributed in the conceptualization of the research problem, study design, write the thesis, analysis of the data, interpretation of the final result.

Aragaw Eshetie Aguade (PhD): Contributed in guidance, consultation, continued follow up, encouragement from the beginning to the end of the study.

Moges Zerihun Fetene: Contributed in guidance, consultation, continued follow up, encouragement from the beginning to the end of the study and formulate the manuscript.

We all authors of the paper carefully read, edited and finally approved the final manuscript.

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List of abbreviations:

ANC – Antenatal Care

AR – Auto-regressive

BP – Blood Pressure

DBP –Diastolic Blood Pressure

PE – Preeclampsia

SBP – Systolic Blood Pressure

REFERENCE

American College of Obstetricians and Gynecologists. (2002). Diagnosis and management of preeclampsia and eclampsia.

American College of Obstetricians and Gynecologists. (2019). Hypertension in pregnancy. Report of the American College of Obstetricians and Gynecologists’ task force on hypertension in pregnancy.

BASCHAT, A. A., DEWBERRY, D., SERAVALLI, V., MILLER, J., BLOCK-ABRAHAM, D., & BLITZER, M. G. (2018). Maternal blood-pressure trends throughout pregnancy and development of pre-eclampsia in women receiving first-trimester aspirin prophylaxis. *Wiley Online Library (wileyonlinelibrary.com)*. DOI: 10.1002/uog.18992.

Belay, A., & Wudad, T. (2019). Prevalence and associated factors of pre-eclampsia among pregnant women attending anti-natal care at Mettu Karl referral hospital, Ethiopia: cross-sectional study. *Clinical hypertension*, 25(1), 14.

Bergen, N., Zhu, G., Asfaw, S., Mamo, A., Abebe, L., Gebretsadik, et al. (2020). Promoting equity in maternal, newborn and child health. Perceptions of public servants in the Ethiopian health sector. *Global Health Action*, 13:1, 1704530, DOI: 10.1080/16549716.2019.1704530.

- Berhan, Y. B. (2014). Causes of maternal mortality in ethiopia: A significant decline in abortion related death. *Ethiop J Health Sci*.
- Biau, D., Kerne, S., & Porcher, R. (2008). The Importance of Sample Size in the Planning and Interpretation of Medical Research. *Clin Orthop Relat Res*, 466:2282–2288. DOI 10.1007/s11999-008-0346-9.
- Bodnar, L., Catov, J., Simhan, H., Holick, M., Powers, R., & Roberts, J. (2007). Maternal vitamin D deficiency increases the risk of preeclampsia. *J Clin Endocrinol Metab*, 92:3517–22.
- Brown, H., & R.Prescott. (2014). Applied mixed models in medicine. *John Wiley & Sons*.
- Bursac, Z., Gauss, C., Williams, D., & Hosmer, D. (2008). Purposeful selection of variables in logistic regression. *Source code for biology and medicine*, 3(1), 17.
- CHEN, L. M., IBRAHIM, J. G., & CHU, H. (2011). Sample size and power determination in joint modeling of longitudinal and survival data. *Statistics in medicine- Wiley Online Library*.
- Chi, Y., & Ibrahim, J. (2006). Joint models for multivariate longitudinal and multivariate survival data. *Biometrics*, 62:432–445. [PubMed: 16918907] .
- Choi, J. (2012). Prediction in the joint modeling of mixed types of multivariate longitudinal outcomes and a time-to-event outcome. *University of Pittsburgh*.
- Choudhury, M., & Friedman, J. (2012). Epigenetics and microRNAs in preeclampsia. *Clinical and experimental hypertension*, 34(5), 334-341.
- Cochran, W. (1946). ‘Relative accuracy of systematic and stratified random samples for a certain class of populations’. *The Annals of Mathematical Statistics*, 17,164–177.
- Duley, L. (2009). The global impact of pre-eclampsia and eclampsia. *Elsevier, perinatology seminar*, 33(3):130-7. 4.
- Endeshaw, M., Abebe, F., Bedimo, M., Asrat, A., Gebeyehu, A., & Keno, A. (2016). Family history of hypertension increases risk of preeclampsia in pregnant women: a case-control study. *Universa Medicina*, 35(3),181-191.
- Faes, C., Aerts, M., Molenberghs, G., Geys, H., Teuns, G., & Bijmens, L. (2008). A high-dimensional joint model for longitudinal outcomes of different nature. *Statistics in medicine*, 27(22), 4408-4427.

- Faucett, C. L., & Thomas, D. C. (1996). Simultaneously modelling censored survival data and repeatedly measured covariates: a Gibbs sampling approach. *Statistics in medicine*, 15, 1663-1685.
- Federal Ministry of Health. (2016). National strategy for newborn and child survival in Ethiopia. Addis Ababa, Ethiopia Maternal and child health directoret, federal ministry of helath.
- Feig, D. S., Ray, J., Lowe, J., Hwee, J., & Booth, G. (2013). Preeclampsia as a risk factor for diabetes: a population-based cohort study. *PLoS medicine*, 10(4).
- Fitzmaurice, G., Molenberghs, G., Davidian, M., & Verbeke, G. (2008). Generalized estimating equations for longitudinal data analysis. In Longitudinal data analysis . *Chapman and Hall/CRC*,51-86.
- Funai, E., Paltiel, O., Malaspina, D., Friedlander, Y., Deutsch, L., & Harlap, S. (2006). Risk factors for pre-eclampsia in nulliparous and parous women:the Jerusalem Perinatal Study. *Paediatric and Perinatal Epidemiology*,19:59–68.
- Gardiner, S., Johnson, C., & Demirel, S. (2012). Factors predicting the rate of functional progression in early and suspected glaucoma. *Investigative ophthalmology & visual science*, 53(7), 3598-3604.
- Garomssa, H., & Dwivedi, A. (2008). Maternal mortality in Ambo Hospital: a five year retrospective review. *Ethiopian J Reprod Health*, 2: 2-13.
- Gaym, A., Bailey, P., Pearson, L., & Admasu, K. (2011). Ethiopian National Em ONCAT. Disease burden due to pre-eclampsia/eclampsia and the Ethiopian health system's response. *International journal of gynaecology and obstetrics: the official organ of the International Federation of Gynaecology and Obstetrics*, 115(1):112-6.
- Ghulmiyyah, L., & Sibai, B. (2012). Maternal mortality from preeclampsia/eclampsia. *In Seminars in perinatology .WB Saunders*, 36(1), 56-59.
- Hickey, G. L., Philipson, P., Jorgensen, A., & Kolamunnage-Dona, R. (2016). Joint modelling of time-to-event and multivariate longitudinal outcomes: recent developments and issues. *BMC medical research methodology*, 16, 117.
- Huang, e. a. (2011). A general joint model for longitudinal measurements and competing risks survival data with heterogeneous random effects. . *Lifetime data analysis*,17(1): 80-100.
- Kang, H. (2013). The prevention and handling of the missing data. *Korean journal of anesthesiology*, 64(5): 402.
- Laird,N.,&Ware,J.(1982). Random-effects models for longitudinal data. *Biometrics*, 38(4):963 -974.

- Little, R., & Rubin, D. (2019). Statistical analysis with missing data. . Vol. 793. John Wiley & Sons.
- Madow, W. G., & Madow, L. H. (1944). 'On the theory of systematic sampling, I'. *The Annals of Mathematical Statistics*,15, 1–24.
- Matter, F., & Sibai, B. (200). Risk factors for maternal morbidity. *Am J Obstet Gynecol*,182(2):307-12.
- Melvin, A., Mohan, K., & Arcuino, L. (1997). Clinical, virologic and immunologic responses of children with advanced human immunodeficiency virus type 1 disease treated with protease inhibitors. *Pediatric Infectious Diseases Journal*, 16:968–974.
- Mo, M., Shen, Y., Si, S., Xin, X., Shao, B., Wang, S., et al. (2019). Feature of trajectory of blood pressure among pregnant women with gestational hypertension. *J Hypertens* ,38:127–132.DOI:10.1097/HJH.0000000000002197.
- Mohammednur, A., & Abdo, N. (2019). Modeling Pregnancy Induced Hypertension among Pregnant Woman in Jimma University Specialized Hospital. *IOSR Journal of Mathematics (IOSR-JM)*e-ISSN: 2278-5728, p-ISSN: 2319-765X. Volume 15, Issue 1 Ser. I (Jan – Feb 2019), PP 22-28.
- Musa, J. M., Ocheke, A., Kahansim, M., Pam, V., & Daru, P. (2018). Incidence and risk factors for preeclampsia inNigeria.*Afri Health Sci*, 18(3):584-595.,<https://dx.doi.org/10.4314/ahs.v18i3.16>.
- Nakimuli, A., Chazara, O., Byamugisha, J., Elliott, A., Kaleebu, P., Mirembe, F., et al. (2014). Pregnancy, parturition and preeclampsia in women of African ancestry. *American journal of obstetrics and gynecology*, 210(6), 510-520.
- Pintilie, M. (2006). Competing Risks: A Practical Perspective. *New York: Wiley*.
- Rizopoulos, D. (2012). Fast fitting of joint models for longitudinal and event time data using a pseudoadaptive gaussian quadrature rule. *Computational Statistics and Data Analysis*,56:491–501.
- Rizopoulos, D. (2012). Joint models for longitudinal and time-to-event data: With applications in R. ed. *CRC Press*.
- Rizopoulos, D. D. (2010). JM: An R package for the joint modelling of longitudinal and time to event data. *Journal of Statistical Software (Online)*, 35(9), 1-33.
- Rocella, E., & Bethesda, M. (2000). Report of the National High Blood Pressure Education Program Working Group on High Blood Pressure in Pregnancy. *Am J Obstet Gynecol*,183:1.

- Rumbold, A. D., Crowther, C., & Haslam, R. (2008). Antioxidants for preventing pre-eclampsia. *Cochrane Database of Systematic Reviews, Issue 1. Art. No.: CD004227. (Systematic Review and Meta-Analysis)* .
- Savita, R., Deepika, P., Anshu, S., & Smiti, N. (2009). Maternal and perinatal outcome in sever preeclampsia and eclampsia. *South Asian Federation of Obstetrics and Gynaecology*, 1: 25-28.
- Schluchter, M. D. (1992). Methods for the analysis of informatively censored longitudinal data. *Statistics in medicine*, 11, 1861-1870.
- Schwarz, G. (1978). Estimating the dimension of a model. *The annals of statistics*, 6(2), 461-464.
- Shamsi, U., Saleem, S., & Nishte, N. (2013). Epidemiology and risk factors of preeclampsia; an overview of observational studies . *Al Ameen J Med Sci*. 6(4):292-300 .*US National Library of Medicine enlisted journal . ISSN 0974-1143* .
- Sibai, B. (2005). Diagnosis, prevention, and management of eclampsia . *Obstet Gynecol*, 105: 402–410.
- Skjærven, R., Wilcox, A., & Lie, R. (2002). The interval between pregnancies and the risk of preeclampsia. *N Engl J Med*, Vol 346, No. 1, [wwwnejm.org](http://www.nejm.org).
- Sousa, H. (2011). Joint Modelling of Longitudinal Quality of Life Measurements and Survival Data in Cancer Clinical Trials. *Doctoral dissertation. Queen's University, Kingston, Canada*.
- Tessema, G., Tekeste, A., & Ayele, T. (2015). Preeclampsia and associated factors among pregnant women attending antenatal care in Dessie referral hospital, Northeast Ethiopia: a hospital-based study. *BMC pregnancy and childbirth*, 15(1), 73.
- Trogstad, L., Magnus, P., Skjaerven, R., & Stoltenberg, C. (2008). Previous abortions and risk of preeclampsia. *International journal of epidemiology*, 37(6):1333-40.
- Verbeke, G., Molenberghs, G., & Rizopoulos, D. (2010). Random effects models for longitudinal data. In *Longitudinal research with latent variables (pp. 37-96)*. Springer, Berlin, Heidelberg.
- Wagnew, M., Dessalegn, M., Worku, A., & Nyagero, J. (2016). Trends of preeclampsia/eclampsia and maternal and neonatal outcomes among women delivering in addis ababa selected government hospitals, Ethiopia: a retrospective cross-sectional study. *The Pan African medical journal*,25(2):12.
- West, B., Welch, K., & Galecki, A. (2014). Linear mixed models: a practical guide using statistical software. *CRC Press*.

- William, C., Shiel, J., FACP, & FACR. (2018). Mediala Definition of systolic. Retrieved from [http://www.medicinenet.com/systolic definition](http://www.medicinenet.com/systolic%20definition).
- World Health Organization. (2011). WHO recommendations for prevention and treatment of pre-eclampsia and eclampsia: evidence base (No. WHO/RHR/11.25). World Health Organization.
- World Health Organization. (2005). World Health Report: Make Every Mother, and Child Count. Geneva: World Health Org.
- Wulfsohn, M., & Tsiatis, A. (1997). A joint model for survival and longitudinal data measured with error. *Biometrics*, 53(1):330–9. doi: 10.2307/2533118.
- Yang, D., Zhu, L., Ran, B., Pu, Y., & Hui, P. (2016). Modeling and analysis of the lane-changing execution in longitudinal direction. *IEEE transactions on intelligent transportation systems*, 17(10), 2984-2992.
- Yanga, L., Yub, M., & Gaoc, S. (2016). Joint Models for Multiple Longitudinal Processes and Time to event Outcome. *J Stat Comput Simul*, 86(18): 3682–3700. doi:10.1080/00949655.2016.1181760.